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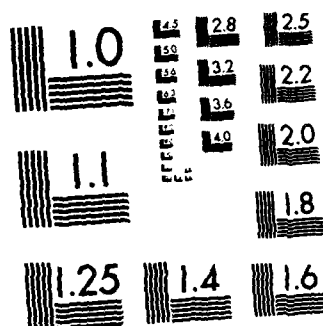
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NAVAL POSTGRADUATE SCHOOL Monterey, California



THESIS

REFINEMENT AND EXTENSION OF SHRINKAGE
TECHNIQUES IN LOSS RATE ESTIMATION OF
MARINE CORPS OFFICER MANPOWER MODELS

by

Charles R. Dickinson

March 1988

Thesis Advisor: R. R. Read

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Refinement and Extension of Shrinkage
Techniques in Loss Rate Estimation of
Marine Corps Officer Manpower Models

by

Charles R. Dickinson
Captain, United States Marine Corps
B.S., United States Naval Academy, 1978

Submitted in partial fulfillment of the
requirements for the degree of

MASTER OF SCIENCE IN OPERATIONS RESEARCH

from the

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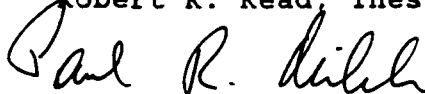


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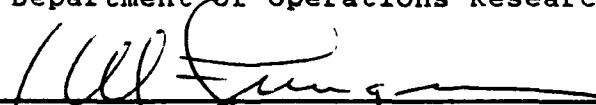
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ABSTRACT

This thesis is a continuation of previous work to apply modern multiparameter estimation techniques to the problem of estimating attrition rates for a large number of small inventory cells in manpower planning models used by the U.S. Marine Corps. The main advances involve the promising introduction of empirical Bayes (non-constant shrinkage) techniques, recognition of the non symmetric nature of the errors with a response to this, and some insight into all aggregation plans that should help provide greater stability for the estimation methods. In addition, the roles of some middle level methodological choices are explored. (Keywords:

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I. INTRODUCTION

A. PURPOSE

In support of project OPUS (Officer Planning and Utility System), Headquarters, U.S. Marine Corps, in 1985, requested assistance in exploring new methods of generating manpower loss rates in order to improve upon the currently used method of force rate analysis. In response, students and faculty in the Operations Research department at the Naval Postgraduate School began exploring the use of modern multiparameter estimation techniques for this problem. Special emphasis is placed on the *small cell* problem, i.e., categories of officer skill, grade and length of service which have low inventory figures. Historically, rate estimators for small cells are unstable and a large number of these cells exist.

This paper builds upon previous work on this problem. The main advances are to append new measures of effectiveness as requested by the Navy Personnel Research and Development Center (NPRDC), introduce a class of empirical Bayes estimators, and explore the effects of some middle level choices in applying the new and existing methods. The ultimate goal is to refine the techniques presented here in order to validate a clear policy for predicting loss rates.

B. BACKGROUND

For an introduction into Marine Corps policy concerning manpower planning, the reader is referred to a thesis submitted in September 1985, at the Naval Postgraduate School, by Major D.D. Tucker [Ref. 1], who presents a detailed background into the Marine Corps officer structure and the manpower planning process. Tucker also provides basic attrition rate theory and calculates overall attrition rates in several different formats as they pertain to the Marine Corps. As the aggregation rate begins to grow smaller with refinement, Tucker illustrates the irregular behavior of losses due to voluntary attrition and the small numbers of losses obtained in such aggregates. Tucker [Ref. 1:p. 50] introduces the James-Stein technique of loss rate estimation as a *shrinking* of individual cell averages toward a grand average in order to reduce the risk, or squared differences of forecast and actual values, the goal being an improvement over the classical maximum likelihood estimator, referred to from now on as the MLE. Tucker chose the ranks of First Lieutenant and Lieutenant Colonel for evaluation using military occupational specialty (MOS) groups of combat support and ground combat (three MOS's were selected for each group). These aggregation schemes were carried through other studies and are included for historical purposes in this paper.

Another thesis submitted at the Naval Postgraduate School in March 1986, by Major J.R. Robinson [Ref. 2] focused on a technique called limited translation of the James-Stein estimator. This particular model attempts to minimize the risk mentioned above by reducing the shrinkage of rate estimates for cells which fall outside a certain range of values centered on the grand mean. It was thought that shrinking all cells by the same rate toward this grand mean may be unwise since the attrition rates for those particular cells, those farthest from the grand mean, may contain attributes which are different than the majority of the cells in the aggregate. Additionally, Robinson's work carries forward the study done by Tucker concerning MLE and James-Stein estimators.

These feasibility studies were limited in scope due to the format of the summary data tape available (acquired from NPRDC), which allows only coarse cell definition. A cell is defined to be a cross-classification of the forty military occupational fields (OF), thirty-one lengths of service (LOS), and ten grades for a total of 12,400 categories for manpower planning purposes [Ref. 2:p. 10]. The data does not distinguish between limited duty officers, those officers specifically designated for limited duty within certain MOS's, and unrestricted officers (regular and reserve) [Ref. 1:p. 20]. These factors dictated a broad aggregation scheme that created stable cell inventories and

allowed the common variance assumption crucial to their work. But some of the application models require more refined information along those lines and experiments with finer aggregation levels, necessary in some real cases, caused this assumption to be violated.

Tucker and Robinson did conclude that the current scheme of attrition rate prediction could be improved upon by their methods, although no dominant scheme was uncovered. Tucker and Robinson did identify problem areas in their studies that required additional work. Tucker [Ref. 1:p. 71] cited a need for a better aggregation method to produce a more homogeneous attrition behavior in all cells. Also, small probabilities of loss within a cell were not dealt with successfully. Robinson [Ref. 2:p. 32] revealed the small cell problem and the inability to normalize the cell means or stabilize the variance with a data transformation. The recommendations of the previous studies highlighted the need for a more refined data tape with current information to include, among others, full Military Occupational Specialty (MOS) information, grade separation to include regular/reserve status, promotion zone data and breakdown of attritions by type.

C. PROGRESS

The project of loss rate estimation, sponsored by NPRDC, is currently moving into the implementation stage. An operational data tape is now available which includes

detailed information on Marine Corps officers for the years 1977-86. Work has begun at the Naval Postgraduate School to break out the data in a useful format and its availability is now nearing completion.

In addition to the results stated earlier, Tucker and Robinson provided loss rate estimates for small cells with no attrition, i.e., MLE equal to zero, and risk effects of the different schemes studied. These results provided the springboard for this study and the requirements for additional future studies.

Captain R.W. Larsen [Ref. 3] has developed a promising aggregation scheme based on cluster analysis, i.e., a classification scheme to aggregate cells which reflect a greater degree of homogeneity in attrition rates than are allowed using current aggregation methods. Table 1 displays the resulting scheme, given in year of current service (YCS) groups proposed for implementation by Larsen.

Table 1. AGGREGATION METHOD PROPOSED BY LARSEN

<u>MOS Category</u>	<u>Bounded YCS Groups</u>
Fixed-Wing Pilots	(1-6,8-19) (7) (20-25) (26)
Rotary-Wing Pilots	(1-5,8-19) (6,7) (20-25) (26)
Naval Flight Officers	(1-5,8-19) (6,7) (20-25) (26)
Lawyers	(1-6,8-19) (7) (20-25) (26)
All Else	(1-3,6-19) (4,5) (20-25) (26)

The parentheses encompass YCS groups which behave similarly. In each case the grade is fixed, i.e., one should aggregate over YCS before aggregating over grade.

This information will allow for more successful application of the loss rate estimation techniques of interest here when the new data tape is provided.

D. GOALS

In an attempt to refine the methods of attrition rate estimation presented in the previous pilot studies, the following goals were set for this paper:

1. Modify the existing methods to:
 - a) use the refined data format,
 - b) extend the work to the case of unequal or non constant variance in the small cell inventory problem,
 - c) study alternate transformation inversion techniques,
 - d) introduce and evaluate additional measures of effectiveness suggested by NPRDC to include, but not limited to, cross validation underages and overages of the deviation between forecasts and actuals,
 - e) examine the graphical effect of shrinkage in the original scale.
2. Introduce empirical Bayes estimation method in several forms for attrition rates for consideration as another option to solve the problem.

II. EXTENSIONS TO ESTIMATION METHODS

A. GENERAL

As stated in Chapter I, Tucker and Robinson published pilot studies introducing the James-Stein, limited translation James-Stein, Maximum Likelihood and Transform Scale Cell Average estimators as alternatives to the aggregate methods previously used for manpower planning in the Marine Corps. Their performance was promising, and this chapter will investigate ways to enhance their performance and possibly develop a dominant estimator for future attrition prediction. Additionally, NPRDC has expressed interest in additional measures of effectiveness for these estimators which may lead to a sharper direction for the goal of producing a viable estimation policy. This chapter concentrates on the following issues concerning previous studies.

B. NON CONSTANT VARIANCE

For a cell having an inventory n , and attrition rate p , the number of attritions y is described by the Binomial (n, p) distribution. The variance of the estimator of $\hat{p} = y/n$ is given by the familiar formula

$$\text{var}(\hat{p}) = \frac{p(1-p)}{n} \quad (2.1)$$

The Freeman-Tukey double arcsine transformation was used by Tucker [Ref. 1:p. 55] and Robinson [Ref. 2:pp. 74-75] in order to map the raw losses to the transformed scale. This formula is given by

$$x = 0.5[n + 0.5]^{1/2} \{ \sin^{-1} [\frac{2y}{(n + 1)} - 1] + \sin^{-1} [\frac{2(y + 1)}{(n + 1)} - 1] \} \quad (2.2)$$

Note that $n=n(t)$ may change with time and thus, so would x . The question of averaging over time before or after applying the Freeman-Tukey transformation is not discussed in this paper. One of the goals of using this transformation is to stabilize the variance of this estimate. In the work of Tucker and Robinson, the assumption of normally distributed random variables with common variance was the setting in which the James-Stein estimator was expected to perform well. Indeed, for n moderate or large and p not too extreme, the variance of the Freeman-Tukey transformation is approximately one. But as discovered by Robinson, this assumption does not hold well when applied to cells with small inventory. A more careful look at the variance of the transformed data is taken here, with particular interest in the region of unstable variance where n and p are small.

Some exploratory graphical work has revealed that the natural log of the variance of x , the transformed value,

behaves quite well as a linear function of $\mu (=E[X])$ and $\mu-1$ in this unstable zone. This leads to the interpolatory formula:

$$\text{var}(x) = 1.6835\mu^{-.8934}(\mu - 1)^{.9881}, \quad (2.3)$$

$$1.001 \leq \mu \leq 2.2$$

The upper limit, $\mu = 2.2$, is the value for which $\text{var}(x) = 1$. Since the least squares fitting process was applied to the logarithms, μ can not be allowed to fall to one or below. The lower limit, $\mu = 1.001$, is an arbitrary value that meets this constraint. The coefficients were validated by a linear regression on the natural log of this formula. The actual variance function for $n = 1, 2, 3, 4, 5, 7, 10$ appear in Figure 1, and the fit is remarkably good for $n \geq 3$. The upper graph shows the region of special interest for this paper and this portion of the curve is fitted by the formula in Equation 2.3 above. The lower plot exhibits the outer tale effect yet to be explored. The small values of $n (< 3)$ do not maintain the stable variance region, i.e., $\text{var}[x]=1$, for long before falling back. As the value for n increases, the more stable the variance becomes. Table 2 is a partial display of the residuals for selected n values, fifteen equally spaced values in the range $1.001 \leq \mu \leq 2.2$, computed by the linear regression formula above. For values of $n > 2$ the difference between actual and fitted values are considered acceptable for this work.

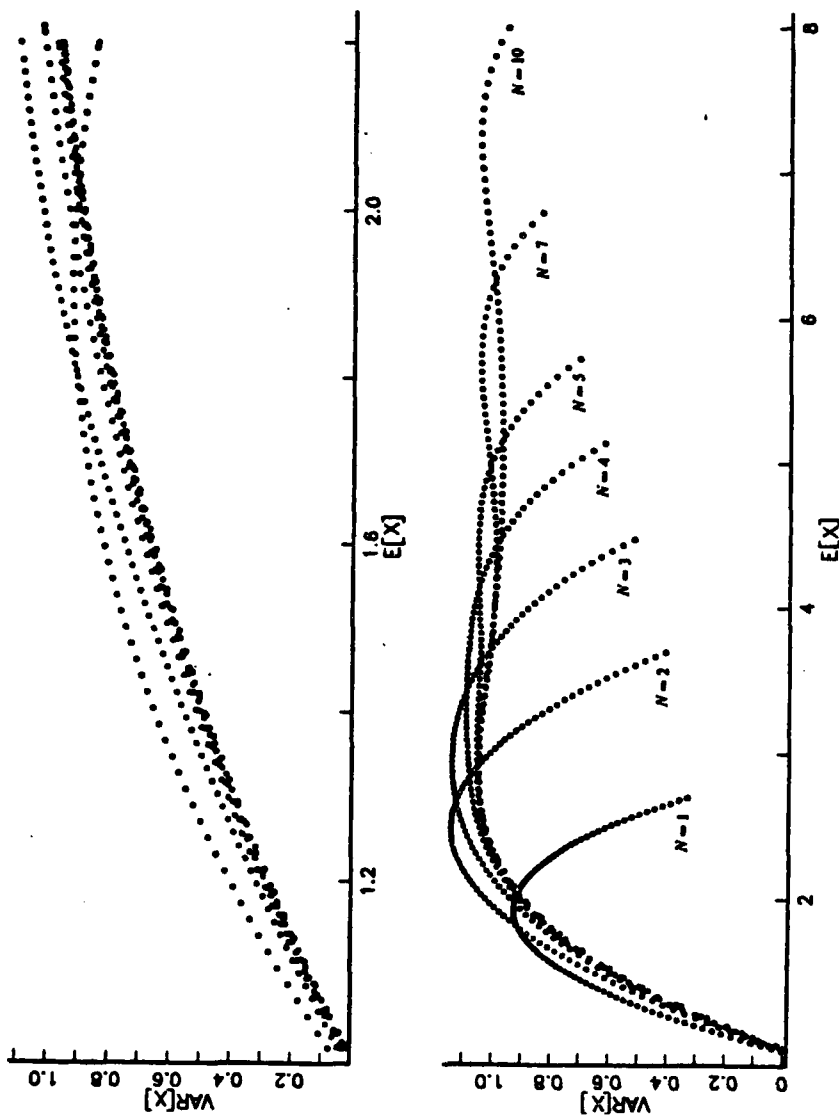


Figure 1. Variance of Freeman-Tukey Transform as Function of It's Mean for Selected n .

Table 2. RESIDUAL LISTING FOR SELECTED VALUES OF N

<u>N=1</u>	<u>N=2</u>	<u>N=4</u>	<u>N=5</u>	<u>N=10</u>
0.07191	0.03912	0.02139	0.017	0.008313
0.08466	0.02569	0.002853	-0.001695	-0.01106
0.1033	0.0239	-0.005448	-0.01049	-0.0206
0.122	0.03002	-0.00655	-0.01185	-0.02257
0.1385	0.04068	-0.004026	-0.009566	-0.02067
0.15	0.05388	0.0000676	-0.005727	-0.0166
0.1552	0.06853	0.004027	-0.002103	-0.01291
0.1527	0.08234	0.007417	0.0009541	-0.009981
0.1414	0.09481	0.00948	0.002557	-0.008299
0.1209	0.1055	0.009825	0.002448	-0.008128
0.08958	0.1129	0.008369	0.0004583	-0.009929
0.04832	0.1169	0.004997	-0.003511	-0.01327
-0.005254	0.1171	0.0001116	-0.00893	-0.0182
-0.06914	0.113	-0.006798	-0.01654	-0.02463
-0.1466	0.104	-0.01534	-0.02585	-0.03366

The boundaries for μ , in the unstable region, as a function of n and p are provided in Figure 2. The formula derived was quite adequate for this study.

C. ALTERNATE TRANSFORM INVERSION METHODS

There are three issues in the inversion of the data from the transformed space. The first is that of selecting values of n , i.e., an average of the $n(t)$'s appearing in the estimation year inventories. This issue is discussed later in the chapter, with the development of a simulation for comparing mean inventory values. The second is that of the inversion of the sum of the two arcsine functions. The third involves the case of time averaging in the transformed scale and the question of what matching form of average

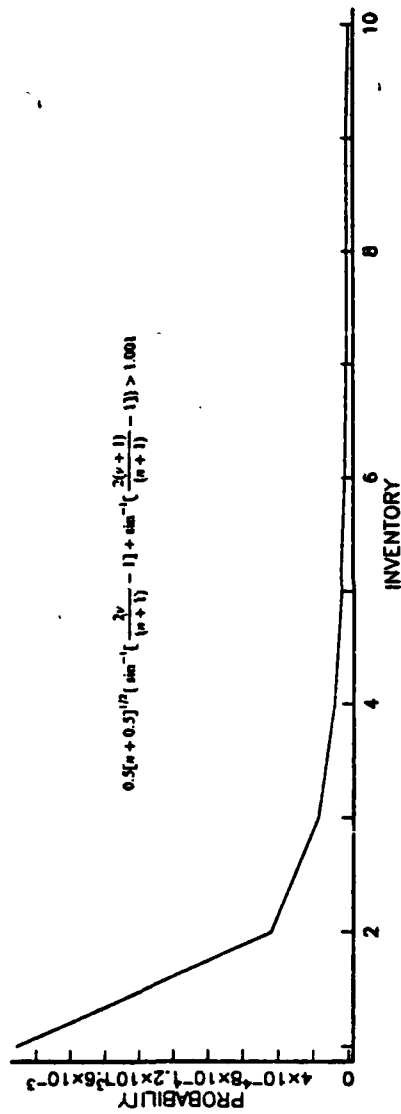
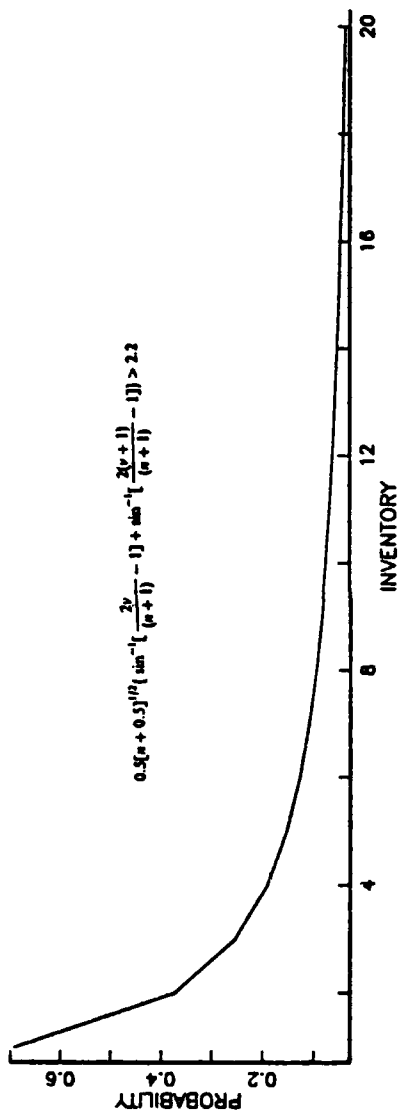


Figure 2. Envelope for the Non-Constant Portion of the Fitted Variance as Functions of Attrition Rate and Inventory.

inventory should be used in the inversion. The third issue is not covered in this paper, and is still pending. Treatment of the second issue will be provided first.

The method used in the pilot studies of Tucker and Robinson, and the most transparent transform inversion formula is:

$$\hat{p}_i = 0.5[1 - \sin(\frac{x}{(n + 1/2)^{1/2}})] \quad (2.4)$$

This will be referred to as the Basic Inversion. The formula offered by Carter and Rolph [Ref. 4] is given by:

$$\hat{p}_i = \sin^2(\hat{\theta}_i) + \left\{ \frac{(1 - 2\sin^2\hat{\theta}_i)}{(4(n/k) + 2)} \right\} \hat{B}_i \quad (2.5)$$

where $\hat{\theta}_i$ is $x/(n + 1/2)^{1/2}$ and \hat{B}_i is the estimated amount of shrinkage in the transformed scale depending on the Bayes model used (constant or proportional prior), k is the number of cells and n is inventory.

Apparently, Carter and Rolph invert prior to shrinking and the above formula provides a shrinkage in the original scale. Rewriting the above, it appears to shrink towards $\hat{p} = 1/2$, and is given by:

$$\hat{p}_i = 1/2 \left(\frac{2\hat{B}_i}{4(n/k) + 2} \right) + \sin^2(\hat{\theta}_i) \left(1 - \frac{2\hat{B}_i}{4(n/k) + 2} \right) \quad (2.6)$$

An exact Freeman-Tukey inversion formula developed by Miller [Ref. 5] was recently uncovered. It is exact in the sense that for a single cell, the formula given by:

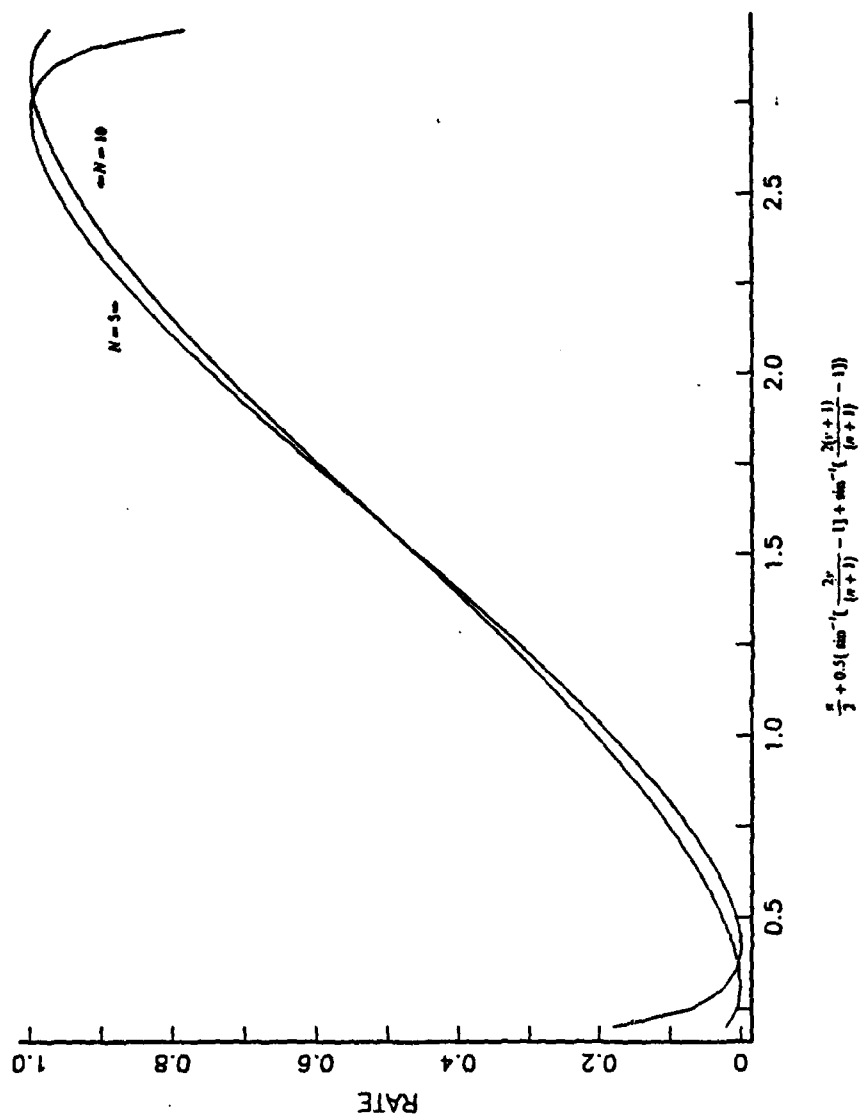
$$\hat{p}(t) = 0.5[1 - \operatorname{sgn}(\cos t)][1 - (\sin t + \frac{\sin 5 - 1/\sin t}{n})^2]^{1/2} \quad (2.7)$$

where $t = x/(n + 1/2)^{1/2} + n/2$ and a is the average central inventory, returns the empirical rate y/n . This will be referred to as the Freeman-Tukey Exact (FTE) inversion. Since this formula is applied using an average value of n , there may be concern that it oscillates between successive integer values of n . Figure 3 indicates, using values of $n=5$ and 10 , that it is quite smooth as an interpolation formula. In the article, Miller suggested using the harmonic mean of the n values provided by the time changes. This issue is examined later in this chapter. The Miller formula was incorporated in the Marine Corps Inventory Attrition Analysis Program provided in Appendix B.

D. SHRINKAGE VISUALIZATION

When shrinkage is a fixed amount for all cells, it is readily visualized, for all degrees of shrinkage, by the linear diagrams presented in the works of Tucker [Ref. 1:p. 52] and Casella [Ref. 6]. Robinson [Ref. 2:p. 20] also presented diagrams of the limited translation option and how the shrinkage is affected. All of these diagrams refer to the transformed scale.

Of more immediate interest is the question of how to visualize shrinkage diagrams of this type in the original scale. Illustrations of diagrams of this type are provided



$$\frac{r}{2} + 0.5 \left(\sin^{-1} \left(\frac{2r}{(r+1)} - 1 \right) + \sin^{-1} \left(\frac{2(r+1)}{(r+1)} - 1 \right) \right)$$

Figure 3. Suitability of the Exact Freeman-Tukey Inversion as an Interpolatory Formula.

using both the basic inversion formula used by both Tucker and Robinson, and the FTE formula. Figure 4 shows how the diagram by Casella [Ref. 6:p. 24] looks in the original scale using the same value of n for all cells. An interesting note to consider is the effect of the difference between inversion formulas, i.e., in the graph which combines the Basic and FTE inverses, the FTE inversion tends to shrink at a slower rate, as values get further from the mean. For a modest transition into the case of non constant n , Figure 5 provides for two values of n . Finally, Figure 6 illustrates the point with many different values of n , taken from actual Marine Corps officer inventory data. In this case, note that Basic inversion shrinks to one original scale value while the FTE inverse is not as focused.

E. CHOICE OF AVERAGE INVENTORY VALUES OVER TIME

A Fortran simulation was developed, included as Appendix C, to test the performance of each of several common means (arithmetic, geometric and harmonic) in the comparison of the inverted estimates with the empirical y/n . Values of n were produced from a Poisson distribution with rate λ ($1 \leq \lambda \leq 20$), and attrition values, y , were produced from the Binomial (n, p) distribution where a was in the range $.01 \leq p \leq .40$. The n and y values were then transformed by using the Freeman-Tukey formula. An exhaustive study of inversion variations with both the FTE formula and the basic formula

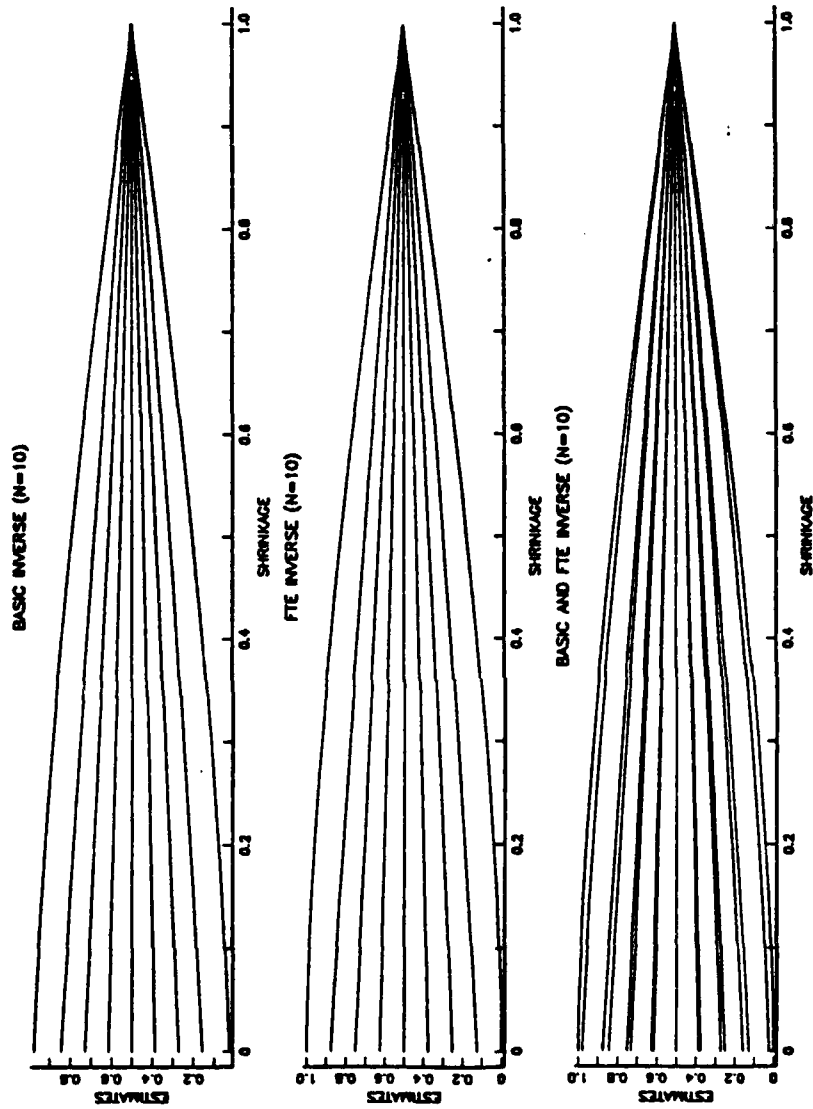


Figure 4. Non Linear Shrinkage Diagrams for a Single Value of Inventory.

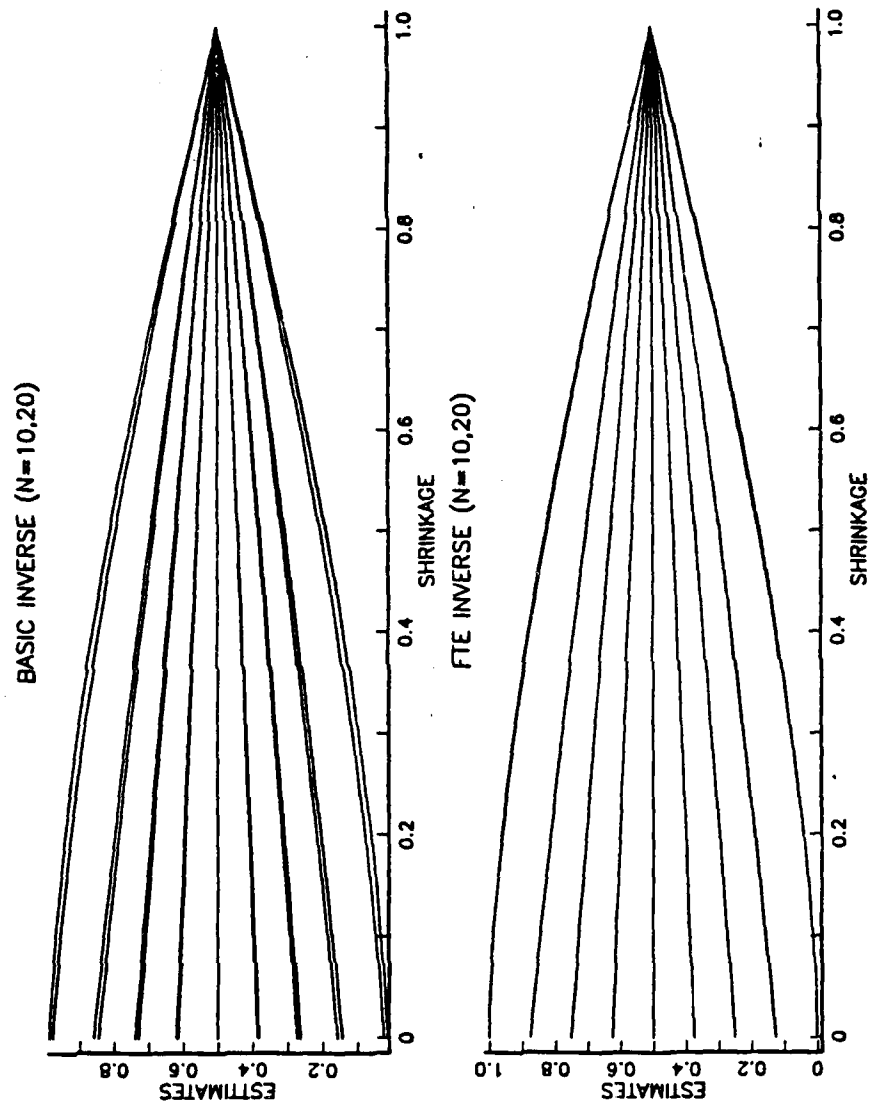


Figure 5. Non Linear Shrinkage Diagrams for Two Values of Inventory.

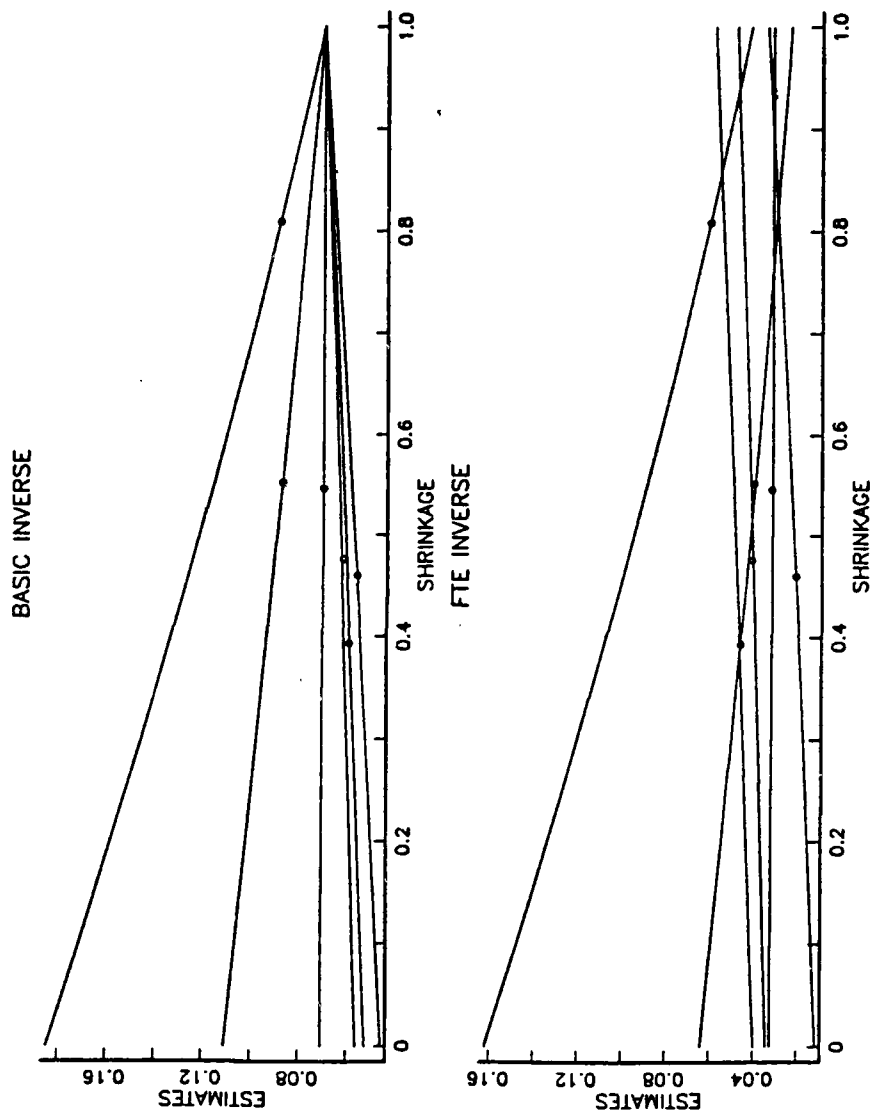


Figure 6. Non Linear Shrinkage Diagrams with Empirical Bayes Estimators for Test Case.

used by Robinson were compared. Table 3 contains selected outputs which are typical of the broader study and shows how the means performed against each other in three different computations listed below.

1. The BASIC Method utilizes the Basic inversion formula and the average transformed scale value.
2. The FTE1 Method utilizes Equation 2.7 with the t values equal to the average (over replications, i.e., time) of $x/(n + 1/2)^{1/2}$ and x given by Equation 2.2. More explicitly, see Equation A.16.
3. The FTE2 Method utilize Equation 2.7 with the t values equal to the average (over replication, i.e., time) of x directly, and x given by Equation 2.2 followed with a division by $(n + 1/2)^{1/2}$

In the context of this study, the arithmetic mean tended to perform as well or better than the geometric or harmonic means. The choice of arithmetic mean also circumvents the problem of an undefined term when $n_1 = 0$ is encountered while using the harmonic or geometric means.

F. QUESTIONS RELATING TO VALIDATION

The pilot studies by Tucker and Robinson both used four years (1977-80) for estimation of the attrition rates and the remaining three years (1981-83) for validation. The estimation rates were used in an attempt to forecast the values for the three validation years. The quality of the forecasts deteriorates as lead time increases in the future. This was anticipated, in fact, Rowe et al., at NPRDC have shown a time series effect in their work forecasting attritions in the Navy [Ref. 7] and are currently

Table 3. SIMULATION OUTPUT OF MEAN VALUES FOR INVERSION UTILIZATION

<u>λ</u>	<u>REPS</u>	<u>P</u>	<u>METHOD</u>	<u>ARITHMETIC</u>	<u>GEOMETRIC</u>	<u>HARMONIC</u>
2.0	40	0.05	BASIC	0.1068	0.1176	0.1308
			FTE1	0.0252	0.0189	0.0122
			FTE2	0.0181	0.0208	0.0245
2.0	40	0.10	BASIC	0.1399	0.1537	0.1708
			FTE1	0.0535	0.0458	0.0369
			FTE2	0.0489	0.0558	0.0648
2.0	40	0.20	BASIC	0.2402	0.2629	0.2906
			FTE1	0.1900	0.1829	0.1739
			FTE2	0.1657	0.1869	0.2144
2.0	40	0.40	BASIC	0.3713	0.4038	0.4429
			FTE1	0.3442	0.3403	0.3353
			FTE2	0.3325	0.3715	0.4213
4.0	40	0.05	BASIC	0.0750	0.0815	0.0921
			FTE1	0.0235	0.0200	0.0147
			FTE2	0.0200	0.0221	0.0258
4.0	40	0.10	BASIC	0.1280	0.1387	0.1564
			FTE1	0.0768	0.0724	0.0655
			FTE2	0.0726	0.0798	0.0922
4.0	40	0.20	BASIC	0.1965	0.2125	0.2386
			FTE1	0.1446	0.1406	0.1339
			FTE2	0.1485	0.1628	0.1869
4.0	40	0.40	BASIC	0.4382	0.4694	0.5185
			FTE1	0.4564	0.4558	0.4549
			FTE2	0.4280	0.4638	0.5223
8.0	40	0.05	BASIC	0.0552	0.0580	0.0616
			FTE1	0.0181	0.0166	0.0148
			FTE2	0.0156	0.0165	0.0177
8.0	40	0.10	BASIC	0.1226	0.1285	0.1363
			FTE1	0.0922	0.0904	0.0880
			FTE2	0.0828	0.0873	0.0933
8.0	40	0.20	BASIC	0.2035	0.2131	0.2255
			FTE1	0.1778	0.1762	0.1741
			FTE2	0.1708	0.1797	0.1915
8.0	40	0.40	BASIC	0.3799	0.3962	0.4170
			FTE1	0.3700	0.3693	0.3684
			FTE2	0.3662	0.3837	0.4065

recommending a second order auto regression scheme for this purpose. In this regard, the work of Tucker and Robinson with validation may be viewed as *forecasting by persistence*, and in light of the NPRDC work, can hardly be expected to perform well in an absolute sense. Their work does, however, provide some comparisons that can be useful in choosing among alternative techniques, and much can be learned from them.

The first measure considered is the Mean Squared Error (MSE) of forecasts. The shrinkage estimation theory provides that these values should be about unity if persistence, independence over time, was agreeable. That is, useful validations can be obtained without modeling the time dependence of the attrition process. Only for a very few cases does this hold (i.e., Warrant Officers, see Tucker). However, if one method of estimation consistently produces noticeably smaller MSE values than do the other methods, this would provide sufficient support for that method's continued development.

The second measure is designed to consider performance in the original scale and is patterned after a chi-square statistic

$$\sum_{i=1}^k \frac{(\text{actuals} - \text{forecasts})^2}{\text{forecasts}} \quad (2.8)$$

If the k cells are independent, then equation 2.8 is approximately a chi-square random variable with k degrees of

freedom. Hence, the mean value is about k and the standard deviation is $(2k)^{1/2}$. Once again, we generally do not see values this small, but use it for comparison purposes. As a side note, the value of k when using the MLE is generally smaller than that of the other estimators due to the omission of those cells with MLE equal to zero over all estimation years, i.e., zero forecasts. This provides an unnatural advantage for the MLE, which must be kept in mind when interpreting the results.

Two disturbing features have emerged from the exploratory analysis. Often a technique that performs well for some cases using the MSE measure is not comparably supported by the chi-square measure. It seems the two measures do not weight the common features equally, and further study is necessary to develop understanding of the results.

Discussion with personnel at NPRDC concerning these matters has led to the introduction of a third measure, which will be called the Mean Absolute Deviation (MAD). This value is computed in the original scale. Moreover, since the cost structure of over-forecasting is not symmetric, we separate the two parts into the average overage deviation given by:

$$ODEV = \frac{1}{k} \sum_{i=1}^k [\text{forecasts} - \text{actuals}]^+ \quad (2.9)$$

and average underage deviation

$$UDEV = \frac{1}{k} \sum_{1}^k [\text{actuals} - \text{forecasts}]^{+} \quad (2.10)$$

where the "+" refers to the positive part, as computed separately. Finally,

$$MAD = ODEV + UDEV \quad (2.11)$$

This measure also has shown that in most cases, the overage exceeds the underage value, which in turn may lead to forecasting the need for an excess number of Marine officers.

III. EMPIRICAL BAYES ESTIMATION

A. BACKGROUND

Although the studies of Tucker and Robinson showed improvement over the current methods, no clear procedure was established and no plans for implementation were drafted. Problem areas were identified while working with small cell inventories and low loss rates. In the transformed scale, the James-Stein assumption of normal distribution of the observations with equal variance did not hold well in such cases. An alternate method of loss rate estimation is required to deal with this special case.

B. INTRODUCTION

Empirical Bayes is simply a generalization of the James-Stein estimator for use with small cell inventories and unstable variance. Random variables are assumed to be of the form:

$$X_i \sim N(\theta_i, d_i), i=1,2,\dots,k \quad (3.1)$$

where θ_i are the unknown parameters to be estimated and d_i are known variances. The Bayesian assumption is that the θ_i 's are themselves from some distribution, in this case also from the normal distribution with mean, ν , and variance, $\rho_1 a$. The proportionality constant ρ_1 is normalized so that

$$\sum \rho_i = k, i=1,2,\dots,k \quad (3.2)$$

where k is the number of observations or cells. Thus,

$$\theta_i \sim N(\nu, \rho_i a), i=1,2,\dots,k. \quad (3.3)$$

This distribution is known as the prior distribution of θ_i and represents the additional assumed information on the value of θ before observing X whose distribution depends on θ [Ref. 8:p. 553]. The use of this information produces an estimate which is a weighted average of the prior mean, ν , and the sample estimate x . In the empirical Bayes context, the values of ν and a , are also estimated from the data since all the information necessary is contained in the marginal distribution of X_i (unconditional on θ_i) [Ref. 6:p. 83]. This chapter focuses on that method.

For a more detailed explanation of the Bayesian family of estimators, the reader is referred to Casella [Ref. 6], Carter and Rolph [Ref. 4] and Efron and Morris [Ref. 9].

C. THE ESTIMATION METHOD

Fay and Herriot [Ref. 10] discuss the empirical Bayes technique and provide an estimate for the prior mean as

$$\nu^* = \frac{\sum_i x_i / (a^* + d_i)}{\sum_i 1 / (a^* + d_i)}, i=1,2,\dots,k \quad (3.4)$$

where a^* is an unbiased estimate of a . The empirical Bayes estimate of θ , commonly called the mean of the posterior

where a^* is an unbiased estimate of a . The empirical Bayes estimate of θ , commonly called the mean of the posterior distribution becomes

$$\delta^{EB} = \left(\frac{a^*}{a^* + d_i} \right) x_i + \left(\frac{d_i}{a^* + d_i} \right) v^*, \quad i=1,2,\dots,k \quad (3.5)$$

This is the mean of the conditional distribution of θ given x [Ref. 8:p. 556]. The weights of each of the prior and sample means are determined by the prior and observed variances.

As a measure of the worth of this estimator, when the number of x 's is ≥ 4 [Ref. 11], the risk, or sum of squared differences between the true unknown parameter and the empirical Bayes estimate, is less than that using the observed x 's, for all θ_i . Efron and Morris [Ref. 9] provide a rigorous proof.

An important result of the empirical Bayes estimator is that shrinkage values are no longer constant, but depend upon information contained in each cell. Additionally, the point toward which the cell values are shrunk is the weighted mean, not the grand mean of the observations. Appendix A provides the derivation of the empirical Bayes estimator, variable shrinkage factors and calculation of the prior variance used in this study as they pertain to the Marine Corps officer attrition rates.

IV. VALIDATION PROCEDURES AND RESULTS

A. GENERAL

The outputs displayed in this chapter are excerpts from that produced by the Marine Corps Personnel Inventory Attrition Analysis (MCIAAP) (see Appendix B). Initially, comparisons will be made with the output displayed by Robinson to identify the validity of the Fortran code with the APL code used by Robinson, then new approaches will be displayed in an attempt to improve the results. The Fortran code in the MCIAAP incorporates most of the procedures studied by Tucker and Robinson. Specifically, the MLE, Transformed Scale Cell Average (TSCA) and the James-Stein (JS) estimators, shrinkage procedures and a basic inversion formula. The FTE inversion formula, the new measures of effectiveness from NPRDC, are also included. For the MAD measure of effectiveness, only maximum likelihood and the modified empirical Bayes variables will be presented, since these showed the most promise over all. All other variables are displayed by the MCIAAP if interest is warranted by the reader. Chapter II describes all of these methods in some detail.

B. HISTORICAL COMPARISONS

Tables 4-9 were developed using the same aggregation scheme and format as Robinson [Ref. 2:pp. 39-41], although

not all of the variables are included (MLE, TSCA, and James Stein are the only variables duplicated in this paper). Additionally, Robinson presented the number of cells, as k values, used for evaluation of each estimate in the aggregate Figure of Merit (FOM) display. These k values are not repeated in the current FOM tables, but are included, as parenthetical values, and discussed in the new MAD measure of effectiveness tables. Since this is the case here, the reader is reminded that the apparent competitive performance of the MLE in the following cases is due, in part, to the possible advantage of having a smaller number of cells used in this MOE's computation.

Since the computations were identical for both studies for these tables, the output was expected to be nearly identical or some pattern discernable between the two. In very few cases was this found to be true, and after verifying that the algorithms coded both in the APL and Fortran versions were correct, we have no reasonable explanation for the discrepancies of the programs. It is suspected that the transfer of raw data into an APL workspace, accomplished by Tucker and Robinson, may be at variance with that into Fortran files, done by contractor.

1. Aviation Aggregate Case

In the case of aviation FOM's, displayed in Table 4, 1981 values for LtCol's in the transformed scale are identical to Robinson's [Ref. 2:p. 39] while the last two

Table 4. AVIATION FIGURES OF MERIT

	1981	1982	1983
TRANSFORMED FOM			
1st Lt			
MLE	4.0863	10.6590	11.1338
TSCA	3.7420	10.3636	10.8711
JS	3.9694	10.5671	11.1626
MOD JS	4.7352	11.2615	12.1206
EB	4.0753	10.7015	11.3303
MOD EB	3.1833	5.0568	6.2870
Lt Col			
MLE	4.3664	8.5033	8.1506
TSCA	5.7768	10.2994	10.3353
JS	5.7369	10.2468	10.2972
MOD JS	5.7208	10.1865	10.2756
EB	5.8482	10.2298	10.3534
MOD EB	4.3301	3.9824	2.7186
ORIGINAL FOM			
1st Lt			
MLE	21.0501	43.1546	53.0538
TSCA	56.4603	93.7810	98.4715
TSCAM	53.7884	90.5164	95.3329
JS	58.0150	95.2150	100.1664
JSM	55.3844	91.9406	97.0163
MOD JS	63.0345	99.8419	105.4897
MOD JSM	60.6413	96.5615	102.3172
EB	58.8824	96.2449	101.2482
EBM	56.2111	92.9699	98.0958
MOD EB	19.7683	38.4637	49.7393
MOD EBM	20.7987	36.3673	48.1659
Lt Col			
MLE	33.9315	59.4200	22.1156
TSCA	38.4478	57.0987	69.0641
TSCAM	35.0390	48.9374	59.6957
JS	37.8248	55.1913	67.1696
JSM	33.2016	46.9901	56.7135
MOD JS	38.0291	51.5325	63.3908
MOD JSM	34.5466	45.1554	52.8285
EB	40.3421	54.9389	67.7091
EBM	36.7498	46.9657	57.8567
MOD EB	40.2056	70.5381	60.5749
MOD EBM	185.7476	65.1553	168.3976

years are smaller. The 1stLt FOM's are larger for all three years. The modified empirical Bayes (Mod EB) values were best overall yet still may be too large (ideal values tend to be close to unity). It is interesting to note that in most cases, the risk values tend to increase with future forecasts, yet the Mod EB FOM decreases as future forecasts increase with time.

In the original scale, MLE tended to be smaller than Robinson's for all years, while TSCA and James-Stein (JS) were all larger. The FTE inversion worked best for all variables except for Mod EB in the first and third validation years. MLE and Mod EB worked best for 1stLt's and are approximately within the range of one standard deviation of the number of cells in the computation. For LtCol's, FTE inverted JS was best with MLE a competitive second, yet all were not in the desired range.

In the MAD measure, Table 5, results are mixed between future forecasts (no set pattern) and variables for the two ranks (FTE Mod EB was best for 1stLt's and MLE was best LtCol's). The mix in overage and underage is most desirable in the first year for all cases, yet the number of cells (k) forecasted as overage exceed the underage forecast cell numbers in most cases. This is cause for concern and the performance of these estimates in this aggregate is suspect. The second year values for LtCol estimates displays a preferred ratio of k values, but the total MOE value may still be too large.

Table 5. MEAN ABSOLUTE DEVIATION (AVIATION)

	1981	1982	1983
1st Lt			
<u>MLE</u>			
UNDERAGE	0.1256(3)	0.0439(1)	0.3069(1)
OVERAGE	2.3985(8)	3.9267(10)	4.2522(10)
TOTAL	2.5240(11)	3.9706(11)	4.5591(11)
<u>MOD EB</u>			
UNDERAGE	0.1822(4)	0.0410(1)	0.3281(1)
OVERAGE	2.3625(8)	3.9040(11)	4.2351(11)
TOTAL	2.5447(12)	3.9451(12)	4.5633(12)
<u>MOD EBM</u>			
UNDERAGE	0.2518(5)	0.0580(2)	0.3420(1)
OVERAGE	2.1654(7)	3.6427(10)	3.9677(11)
TOTAL	2.4172(12)	3.7007(12)	4.3098(12)
Lt Col			
<u>MLE</u>			
UNDERAGE	0.6005(5)	1.5139(10)	0.0396(1)
OVERAGE	0.7925(8)	0.3928(3)	1.3467(12)
TOTAL	1.3930(13)	1.9067(13)	1.3863(13)
<u>MOD EB</u>			
UNDERAGE	0.6053(6)	1.4717(9)	0.0841(2)
OVERAGE	1.3083(10)	1.0970(8)	2.2331(14)
TOTAL	1.9136(16)	2.5687(17)	2.3173(16)
<u>MOD EBM</u>			
UNDERAGE	0.6890(6)	1.5366(10)	0.1167(2)
OVERAGE	1.0120(10)	0.7179(7)	1.7501(14)
TOTAL	1.7010(16)	2.2546(17)	1.8669(16)

2. Combat Support Aggregate Case

Table 6 displays the combat support aggregate output from the program. In the transformed scale comparison with Robinson [Ref. 2:p. 40], the values for the two ranks take opposite directions. Where the 1stLt FOM values are all

Table 6. COMBAT SUPPORT FIGURES OF MERIT

	1981	1982	1983
	TRANSFORMED FOM		
1st Lt			
MLE	2.1877	2.9028	2.2447
TSCA	1.6329	2.4580	1.8696
JS	1.6925	2.4816	1.8987
MOD JS	2.3684	3.0520	2.5356
EB	2.3323	3.1991	2.6868
MOD EB	2.2293	2.0048	2.1531
Lt Col			
MLE	1.3103	1.6978	2.3108
TSCA	0.7502	1.4070	1.7540
JS	0.6421	1.2345	1.5429
MOD JS	0.7009	1.0929	1.3082
EB	0.7736	1.2589	1.5799
MOD EB	1.8691	2.1236	4.9017
	ORIGINAL FOM		
1st Lt			
MLE	111.4260	112.1917	70.2767
TSCA	73.5008	90.0523	73.9380
TSCAM	79.0947	93.8244	80.9764
JS	76.9945	89.8064	71.5890
JSM	429.8503	84.4175	65.2639
MOD JS	143.1700	113.0083	90.6783
MOD JSM	168.9385	114.7083	92.6599
EB	102.9525	119.0900	96.3373
EBM	108.3578	111.2701	86.8138
MOD EB	68.7251	63.9054	48.2719
MOD EBM	176.3961	59.5760	43.0037
Lt Col			
MLE	36.0152	66.0261	42.4097
TSCA	61.7769	52.0346	99.7917
TSCAM	65.9535	49.6307	107.1016
JS	61.0511	49.2280	96.7861
JSM	66.3509	47.8490	118.2564
MOD JS	66.5470	49.9366	99.2566
MOD JSM	81.4722	50.2637	103.8460
EB	69.7508	55.9812	104.5311
EBM	71.5728	50.0281	109.2012
MOD EB	62.6316	50.3125	96.8693
MOD EBM	83.1337	48.6576	127.1541

larger (except for MLE in 1981), the LtCol values are smaller. Overall, all estimates are competitive with each other over all validation years with TSCA performing best for the 1stLt's and James-Stein outperforming the rest in the LtCol's. Note here also that the Mod EB estimate did not fair well in either case, and was the worst estimate in risk for the LtCol's.

In the original scale, a different and more random effect is noticed between Table 6 and the results of Robinson. Risk values tended to fluctuate with forecast lead time between estimates for 1stLt's where the LtCol values were more stable, yet were quite different from the transformed scale pattern. In this case, Mod EB performed well for 1stLt's, and was the only one within the desired range of standard deviations. MLE for the LtCol's was also the only estimate desirable for this measure. The Miller inverse showed poorly in this aggregate for all estimation methods, in every case exceeding it's counterpart.

In the MAD measure of effectiveness, Table 7, MLE performed the best in both ranks. The Mod EB effect is competitive in the 1stLt's case, and all variables in the LtCol case are closely grouped. The number of cells are heavily weighted towards an overage forecast in all three validation years, throughout this aggregate. Once again, this is not a desirable quality for a MOE.

Table 7. MEAN ABSOLUTE DEVIATION (COMBAT SUPPORT)

	1981	1982	1983
1st Lt			
<u>MLE</u>			
UNDERAGE	0.3737(14)	0.1617(8)	0.1781(8)
OVERAGE	0.6731(19)	0.9574(25)	0.8835(25)
TOTAL	1.0468(33)	1.1191(33)	1.0616(33)
<u>MOD EB</u>			
UNDERAGE	0.4028(15)	0.1275(9)	0.1427(9)
OVERAGE	0.7311(40)	0.9972(45)	0.9422(38)
TOTAL	1.1339(55)	1.1246(54)	1.0849(47)
<u>MOD EBM</u>			
UNDERAGE	0.4592(16)	0.1603(9)	0.1856(10)
OVERAGE	0.5930(39)	0.8318(45)	0.7575(37)
TOTAL	1.0522(55)	0.9921(54)	0.9431(47)
Lt Col			
<u>MLE</u>			
UNDERAGE	0.1360(9)	0.1688(7)	0.0541(5)
OVERAGE	0.1510(14)	0.2591(16)	0.3415(18)
TOTAL	0.2870(23)	0.4280(23)	0.3957(23)
<u>MOD EB</u>			
UNDERAGE	0.1638(11)	0.1721(7)	0.1162(6)
OVERAGE	0.2232(29)	0.3606(35)	0.4548(29)
TOTAL	0.3870(40)	0.5327(42)	0.5710(35)
<u>MOD EBM</u>			
UNDERAGE	0.2044(11)	0.1901(7)	0.1408(6)
OVERAGE	0.1263(29)	0.2240(35)	0.2589(29)
TOTAL	0.3307(40)	0.4142(42)	0.3996(35)

3. Ground Combat Aggregate Case

In the final aggregate studied by Robinson [Ref. 2:p. 41], that of ground combat MOS's, the results again, as displayed in Table 8, were mixed between computational methods. In the original scale, value differences tended to be larger than Robinson's in all cases for 1stLt's, showing a proportional growth over time. Conversely, LtCol values were smaller in the first two validation years and slightly larger in the final year. TSCA looks good for the 1stLt's with the others being competitive, yet the size of the values may not be desirable. Mod EB turned out the best for LtCol's, but all other estimates are competitive and in the desired range.

On the original scale, all values for the estimates for 1stLt's are too large, with Mod EB performing well in that regard. LtCol values are all larger than Robinson's, but the estimates are all within range of one standard deviation and, therefore, competitive. MLE, once again, shows well for this rank. In opposing fashion to that of the combat support aggregate, the FTE inversion outperformed all Basic inversions for all variables.

The MAD measure of effectiveness, Table 9 confirms the results found above in that Mod EB was best for 1stLt's and MLE did well for LtCol's. The values for 1stLt were found to be too large overall, while all estimates for LtCol were competitive and in the desired range. The mix between

Table 8. GROUND COMBAT FIGURES OF MERIT

	1981	1982	1983
TRANSFORMED FOM			
1st Lt			
MLE	2.5764	4.4136	6.8744
TSCA	1.9741	3.8072	6.1152
JS	2.1737	3.9738	6.3970
MOD JS	3.0859	4.7919	7.5619
EB	2.5558	4.3626	6.9607
MOD EB	3.1538	3.5202	4.7069
Lt Col			
MLE	1.2126	1.9674	3.1084
TSCA	1.5715	2.3553	3.5566
JS	1.5571	2.2719	3.4365
MOD JS	1.6227	2.1317	3.1883
EB	1.7087	2.2464	3.3345
MOD EB	1.1964	2.5830	2.5029
ORIGINAL FOM			
1st Lt			
MLE	87.3362	119.2637	143.9128
TSCA	87.9712	116.6767	145.6236
TSCAM	83.7648	111.3995	142.0902
JS	95.4717	122.9425	152.7354
JSM	92.0238	117.5264	149.3965
MOD JS	124.6490	147.6611	180.5502
MOD JSM	134.9472	142.5461	178.6173
EB	109.6349	136.6321	167.3234
EBM	105.5708	131.0400	163.8189
MOD EB	64.3647	81.2071	101.9511
MOD EBM	63.6623	77.8159	100.7150
Lt Col			
MLE	35.0538	45.8172	53.0111
TSCA	55.9649	198.8659	123.1306
TSCAM	43.5234	201.9953	122.8648
JS	56.3346	198.9037	122.2786
JSM	43.9957	201.3299	115.7143
MOD JS	60.0730	202.2959	122.9959
MOD JSM	52.3940	210.7551	121.4254
EB	62.7441	205.4249	126.6544
EBM	50.4049	209.3853	125.7642
MOD EB	48.0871	197.7316	111.9375
MOD EBM	38.1231	323.5654	242.4398

over-estimates and under-estimates, i.e., the number of cells forecasted over or under the actuals, is not a desirable one, yet the same mix holds true for all validation years. This also was the finding of both preceding aggregates.

Table 9. MEAN ABSOLUTE DEVIATION (GROUND COMBAT)

	1981	1982	1983
1st Lt			
<u>MLE</u>			
UNDERAGE	0.1525(5)	0.1855(10)	0.2466(10)
OVERAGE	1.9322(30)	2.5979(25)	3.1691(25)
TOTAL	2.0847(35)	2.7834(35)	3.4157(35)
<u>MOD EB</u>			
UNDERAGE	0.2340(10)	0.1323(8)	0.2220(8)
OVERAGE	1.7743(33)	2.3086(32)	2.8449(30)
TOTAL	2.0083(43)	2.4409(40)	3.0669(38)
<u>MOD EBM</u>			
UNDERAGE	0.2618(10)	0.1705(10)	0.2617(10)
OVERAGE	1.6492(33)	2.1738(30)	2.6935(28)
TOTAL	1.9111(43)	2.3443(40)	2.9552(38)
Lt Col			
<u>MLE</u>			
UNDERAGE	0.2313(10)	0.3034(11)	0.1368(10)
OVERAGE	0.4182(25)	0.4202(24)	0.6453(25)
TOTAL	0.6495(35)	0.7236(35)	0.7821(35)
<u>MOD EB</u>			
UNDERAGE	0.2315(11)	0.3299(14)	0.1426(12)
OVERAGE	0.6392(34)	0.6922(32)	0.9523(31)
TOTAL	0.8706(45)	1.0221(46)	1.0950(43)
<u>MOD EBM</u>			
UNDERAGE	0.2650(12)	0.3790(15)	0.1864(12)
OVERAGE	0.4071(33)	0.4427(31)	0.6535(31)
TOTAL	0.6720(45)	0.8217(46)	0.8399(43)

4. Test Aggregate Case

In an effort to experiment with aggregate groupings, another case was built upon which all estimation methods were applied, and the results compared with previously studied groups. The grade was held fixed, in this case the grade of Captain was used, with a LOS in the eighth year. Several MOS's were chosen so that to insure an aggregate with non-zero inventories. In the transformed scale, Table 10, all estimates have small and competitive values, with Mod EB and MLE doing well. In the original scale, once again, all estimates are well within range, but Mod EB and MLE were the standouts. The FTE inversion option did not fair well with this aggregate. All estimates in the MAD measure of effectiveness, Table 11, are very competitive, and the first validation year mix (overage and underage) may be the best yet. It is interesting to note that the cell split in the first validation year between overage and underage forecasts is contrary to that of the historical displays above. Although the total number of cells used ($k=6$) is much smaller than the previous table, it is felt that the aggregate scheme was the major factor in this finding, and the results are promising.

Table 10. TEST AGGREGATE FIGURES OF MERIT

	1981	1982	1983
TRANSFORMED FOM			
Captain			
MLE	0.9603	2.3463	2.0842
TSCA	1.4151	3.3033	2.9152
JS	1.3753	3.3969	2.8725
MOD JS	1.3943	3.8165	2.8829
EB	1.6840	4.2698	3.2957
MOD EB	0.8249	1.7132	1.2980
ORIGINAL FOM			
Captain			
MLE	3.0402	4.8104	4.9399
TSCA	4.6161	10.2779	7.5714
TSCAM	8.1467	7.6241	8.8347
JS	4.3970	10.3000	7.4726
JSM	7.3199	7.6666	8.3860
MOD JS	4.2990	10.7640	7.6371
MOD JSM	6.0079	8.3070	7.9259
EB	5.3660	12.0337	8.8679
EBM	6.7181	9.3915	8.7959
MOD EB	1.4463	6.0323	3.9066
MOD EBM	4.6465	4.3354	5.9175

Table 11. MEAN ABSOLUTE DEVIATION (TEST CASE)

	1981	1982	1983
Captain			
<u>MLE</u>			
UNDERAGE	0.2625(3)	0.0463(1)	0.2097(2)
OVERAGE	0.1351(3)	0.7374(5)	0.4391(4)
TOTAL	0.3976(6)	0.7837(6)	0.6488(6)
<u>MOD EB</u>			
UNDERAGE	0.1342(4)	0.0000(0)	0.0965(1)
OVERAGE	0.2575(2)	0.9601(6)	0.6306(5)
TOTAL	0.3918(6)	0.9601(6)	0.7270(6)
<u>MOD EBM</u>			
UNDERAGE	0.3712(4)	0.0809(1)	0.2352(2)
OVERAGE	0.0852(2)	0.6148(5)	0.3025(4)
TOTAL	0.4563(6)	0.6957(6)	0.5377(6)

V. CONCLUSIONS AND RECOMMENDATIONS

A. RESULTS

Omitting the fact of differences in computational methods between this paper and the Robinson thesis, it is clear that the results are still inconclusive at this time. In the historical comparisons, the estimates studied varied in their performance from one aggregate to another. The different risk calculations, in the transformed and original scales, as well as the different measures of effectiveness showed little agreement in estimation methods. Long range estimation still remains a problem since with time, the estimate risks increase significantly. Much emphasis was put on the modified empirical Bayes estimator in this paper and those specific results look promising under possibly different and more stable aggregate conditions.

One area of mention here concerning the results has to do with the composition of inventories used in the selected aggregates. It was noticed that the range of inventory (smallest to largest value in any particular MOS) was quite large, and may be affecting the results. In fact, Carter and Rolph [Ref. 4:p. 883] managed this point by further disaggregating into levels of activity as it concerned their work with fire alarm estimation.

B. CONCLUSIONS

The results of the previous feasibility studies and this paper leads to several conclusions. The chosen aggregate still seems to be the key element in the work of attrition estimation. It does seem clear that the aggregation scheme should be focused on fixed ranks, in narrow LOS ranges and multiple MOS's. This conclusion is based upon the relative success of the test aggregate utilized in this work, and upon similar success by Carter and Rolph. Additionally, the partitioning of inventory levels in the aggregate should stabilize the estimation method performance. All the estimation method studied to date have their strengths, but still, no one method of estimation excels over the others.

C. RECOMMENDATIONS

No estimation method studied in this paper is recommended for implementation at this time. In addition, no method is recommended for removal from consideration since basic aggregation problems are felt to be the largest contributory factor to the mixed results found here. These recommendations are presented for further study.

1. The aggregation method offered by Captain Larsen [Ref. 3] should be examined further with its relevance to the scheme in the test case used here.
2. Utilization of the detailed data provided by NPRDC as soon as possible in order to expand the available information and open new avenues of aggregation.
3. Re-examine the inventory mean value used in the FTE inverse to determine its future role in the estimation process.

APPENDIX A. ESTIMATION ALGORITHMS

A. NOTATION

The following notation is a generalization of that used by Robinson [Ref. 2:p. 71] as it pertains to the aggregation scheme used in this paper. The indexing system has been changed to reflect the use of officer grades to identify the cells.

i = index of MOS cells in the aggregate.

j = index of LOS cells in the aggregate.

k = index of Grade cells in the aggregate.

t = index of time periods in the aggregate.

$\text{inv}(i,j,k,t)$ = inventory with MOS i , LOS j and Grade k during year t .

$y(i,j,k,t)$ = number of attritions in cell (i,j,k) during year t .

$n(i,j,k,t)$ = maximum $\{y(i,j,k,t), .5[\text{inv}(i,j,k,t) + \text{inv}(i,j,k,t + 1)]\}$

D = matrix identifying cells with no inventory over all estimation years.

$d(i,j,k) = 0$, if cell is a structural zero.

$d(i,j,k) = 1$, if cell is not a structural zero.

B. FREEMAN-TUKEY EXACT INVERSION ALGORITHM

This algorithm provides a formula for the inverse of the Freeman-Tukey double arcsine transformation to express the means of the transformed values into proportions on the original scale.

STEP 1. Transform the data using the Freeman-Tukey double arcsine transformation.

$$x(i,j,k,t) = 0.5[n(i,j,k,t)+0.5]^{1/2} \left\{ \sin^{-1} \left[\frac{2(y(i,j,k,t))}{(n(i,j,k,t)+1)} - 1 \right] + \sin^{-1} \left[\frac{2(y(i,j,k,t)+1)}{(n(i,j,k,t)+1)} - 1 \right] \right\} \quad (A.1)$$

STEP 2. Calculate the average cell inventory and transformed values over all T.

$$\bar{x}(i,j,k) = (1/T) \sum_t x(i,j,k), \quad (A.2)$$

$$\overline{\text{inv}}(i,j,k) = (1/T) \sum_t \text{inv}(i,j,k). \quad (A.3)$$

STEP 3. Check for inversion values less than zero or greater than one. The algorithm is terminated at this point if these conditions are met.

$$tm = \bar{x}(i,j,k) / (\overline{\text{inv}}(i,j,k) + 0.5)^{1/2} + \pi/2, \quad (A.4)$$

$$tm_1 = \sin^{-1} (1 / (\overline{\text{inv}}(i,j,k) + 1))^{1/2}, \quad (A.5)$$

If $tm < tm_1$, FTE = 0.0,

IF $tm > (\pi - tm_1)$, FTE = 1.0,

STEP 4. Compute the sgn function.

If $\cos(tm) > 0.0$ $sgn = 1.0$,

If $\cos(tm) < 0.0$ $sgn = -1.0$.

STEP 5. Computer the inverted value of the transformed variable

$$FTE = 0.5(1 - sgn[1 - [\sin(tm) + \frac{\sin(tm) - 1/\sin(tm)}{inv(i,j,k)}]^2]^{1/2}) \quad (A.6)$$

C. TRANSFORM SCALE VARIANCE ALGORITHM

This algorithm produces the variance of the transformed data within the unstable variance range where the assumption of normality does not hold. As discussed in Chapter II, this range of instability was determined to be between 1.001 and 2.2. The method for computing the variance within this range is presented here.

STEP 1. Compute the transformed scale values to be used for the variance. Since $x(i,j,k)$ was computed in the FTE inversion algorithm of this appendix, the formulae will not be repeated.

$$\bar{x}(t) = \sum_{ijk} [0.5 + inv(i,j,k,t)]^{1/2}, \quad (A.7)$$

$$z(i,j,k) = \bar{x}(i,j,k) + \frac{x(t)}{T}(\pi/2). \quad (A.8)$$

STEP 2. Compute the variance of transformed cell values. The formula is derived in Chapter II as a linear regression equation. The range of the variance is limited to $0.05 \leq \text{var}(i,j,k) \leq 1.0$.

$$\begin{aligned} \text{var}(i,j,k) &= az(i,j,k)^{b_1} (z(i,j,k) - 1)^{b_2}, \\ 1.001 &\leq z(i,j,k) \leq 2.2 \end{aligned} \quad (\text{A.9})$$

where $a = 1.6835$, $b_1 = -0.8934$, and $b_2 = 0.8991$.

D. EMPIRICAL BAYES SHRINKAGE ALGORITHM

This algorithm computes the empirical Bayes shrinkage values and estimator of attrition rates. Prior to entering this algorithm, the value for a has been set to 0.0.

STEP 1. Initialization.

$$a_i = a. \quad (\text{A.10})$$

STEP 2. Define the α and γ parameters in the empirical Bayes context.

$$\alpha = \sum_{ijk} 1/(a + \text{var}(i,j,k)), \quad (\text{A.11})$$

where $\text{var}(i,j,k)$ is defined by formula (A.9).

$$\gamma = \sum_{ijk} \alpha / \sum_{ijk} 1/(\alpha + \text{var}(i,j,k)). \quad (\text{A.12})$$

STEP 3. Compute the empirical Bayes prior variance.

$$a = a - \frac{k-1 - \sum_{ijk} \alpha [\bar{x}(i,j,k) - \bar{z}]^2}{\sum_{ijk} [\alpha (\bar{x}(i,j,k) - \bar{z})^2]} \quad (A.13)$$

where $z = \sum_{ijk} \gamma \bar{x}(i,j,k)$ and $k = \sum_{ijk} d(i,j,k)$.

STEP 4. Check for the following conditions and branch accordingly.

If, $a < 0.0$ set $a = 0.0$ and go to step 5.

If, $|a - a_i| > .0001$ return to Step 1.

STEP 5. Compute the empirical Bayes estimator.

$$\hat{x}^{EB}(i,j,k) = \left[\frac{a}{a + \text{var}(i,j,k)} \right] \bar{x}(i,j,k) + \left[\frac{\text{var}(i,j,k)}{a + \text{var}(i,j,k)} \right] \bar{z}. \quad (A.14)$$

STEP 6. Compute the empirical Bayes shrinkage set.

$$\text{shr}^{EB}(i,j,k) = \text{var}(i,j,k) / (a + \text{var}(i,j,k)). \quad (A.15)$$

E. MODIFIED EMPIRICAL BAYES SHRINKAGE ALGORITHM

This algorithm differs only slightly from that of the empirical Bayes shrinkage algorithm listed above, and for

that reason, only the pertinent steps will be displayed here.

Consider a particular cell and its transformed value $x(t)$ for $t=1, \dots, T$ in the estimation set. To form a weighted time average, set

$$x_t(t) = x(t) / [0.5 + \text{inv}(t)]^{1/2} \quad (\text{A.16})$$

This leads to a modification of step 1 in the following form.

STEP 1. Compute the weighted transformed scale average.

$$\overline{x_t} = \frac{1}{T} \sum_t x_t(t) \quad (\text{A.17})$$

Since the weights will modify the variance of the weighted transformed scale values.

$$\text{var}[x_t(t)] = \text{var}(i, j, k) / (0.5 + \text{inv}(t)) \quad (\text{A.18})$$

$$\text{var}[\overline{x_t}(t)] = \left(\frac{1}{T}\right)^2 \sum_t \text{var}[x_t(t)] = v_t / T \quad (\text{A.19})$$

The extreme right of formula A.19 serves to define v_t , the approximation for the weighted variance in the validation set for that cell.

The values of $\overline{x_t}$ and $\text{var}(\overline{x_t})$ are carried throughout this version of empirical Bayes estimation to form the values XTEB for the empirical Bayes attrition rate. For the purposes of inversion to the original scale, the usual Basic

formula is modified since the weighting factor, $0.5 + \text{inv}$, is already incorporated into the transform value. The new Basic inversion for this version is simply as follows:

$$p = 1/2[1 + \sin(XTEB)] \quad (A.20)$$

APPENDIX B. MARINE CORPS PERSONNEL INVENTORY ATTRITION ANALYSIS PROGRAM

PROGRAM MARCOR	MAR00010
*****	MAR00020
* MARINE CORPS PERSONNEL INVENTORY ANALYSIS PROGRAM	* MAR00030
*	* MAR00040
* PURPOSE: TO PROVIDE FUTURE PERSONNEL LOSS ESTIMATION OF THE	* MAR00050
* MARINE CORPS OFFICER STRUCTURE BY SEVERAL DIFFERENT	* MAR00060
* ESTIMATION SCHEMES.	* MAR00070
* PROGRAMMERS: LUIS URIBE, INDEPENDENT CONTRACTOR	* MAR00080
* CAPT C R DICKINSON, USMC	* MAR00090
* DESCRIPTION: WRITTEN IN FORTRAN 77 FOR THE IBM 3033 MAINFRAME	* MAR00100
* COMPUTER RESIDENT AT NPS, THIS PROGRAM INCORPORATES	* MAR00110
* METHODS EXPLORED BY MAJOR D.D. TUCKER , MAJOR J.R.	* MAR00120
* ROBINSON AND CAPTAIN DICKINSON. ESTIMATION SCHEMES	* MAR00130
* INCLUDED ARE:	* MAR00140
* MAXIMUM LIKELIHOOD ESTIMATION (MLE)	* MAR00150
* TRANSFORMED SCALE CELL AVERAGE (TSCA)	* MAR00160
* JAMES-STEIN	* MAR00170
* LIMITED TRANSLATION JAMES-STEIN	* MAR00180
* EMPIRICAL BAYES ESTIMATION	* MAR00190
* RISK ANALYSIS WAS PERFORMED IN BOTH THE ORIGINAL AND	* MAR00200
* TRANSFORMED SCALES. ON THE TRANSFORMED SCALE,	* MAR00210
* ACHIEVED BY USING THE FREEMAN-TUKEY DOUBLE ARCSINE	* MAR00220
* TRANSFORMATION, THE RISK WAS DEFINED BY THE AVERAGE	* MAR00230
* SQUARED DEVIATION OF ACTUAL AND ESTIMATED VALUES.	* MAR00240
* FOR THE ORIGINAL SCALE, THE TRANSFORMED VALUES WERE	* MAR00250
* FIRST INVERTED USING TWO TECHNIQUES. AN AD-HOC	* MAR00260
* FORMULA USED BY TUCKER AND ROBINSON AND AN EXACT	* MAR00270
* INVERSE FORMULA FROM JOHN J MILLER. TWO MEASURES OF	* MAR00280
* EFFECTIVENESS, CHI-SQUARE GOODNESS OF FIT AND MEAN	* MAR00290
* ABSOLUTE DEVIATION (MAD), WERE THEN EVALUATED.	* MAR00300
* RESULTS ARE PRINTED, WITH ALL ARRAYS COMPUTED, AT THE	* MAR00310
* COMPLETION OF THE PROGRAM.	* MAR00320
* VARIABLE DESCRIPTION:	* MAR00330
* THE FOLLOWING LIST OF VARIABLES, WITH A BRIEF	* MAR00340
* DESCRIPTION, FORM THE BASIS FOR THE VARIABLE NAMES	* MAR00350
* FOR THE ENTIRE PROGRAM.	* MAR00360
* MLE - MAXIMUM LIKELIHOOD ESTIMATOR	* MAR00370
* TSCA - TRANSFORMED SCALE CELL AVERAGE	* MAR00380
* J - JAMES-STEIN ESTIMATOR	* MAR00390
* J1 - JAMES-STEIN ESTIMATOR MODIFIED BY	* MAR00400
* A VARIATION ON SHRINKAGE	* MAR00410
* JLT - LIMITED TRANSLATION JAMES-STEIN	* MAR00420
* EB - EMPIRICAL BAYES	* MAR00430
* TEB - EMPIRICAL BAYES MODIFIED BY A	* MAR00440
* VARIATION ON	* MAR00450
* TRANSFORMED SPACE VARIABLES	* MAR00460
* THE FOLLOWING ARE UNIQUE VARIABLES IN THE TRANSFORMED	* MAR00470
* SPACE:	* MAR00480
* X - VALUE OF THE FREEMAN-TUKEY	* MAR00490

*	TRANSFORMATION	*	MAR00500
*	XB - (X-BAR) AVERAGE TRANSFORM VALUE	*	MAR00510
*	OVER ESTIMATION YEARS	*	MAR00520
*	XBB - GRAND MEAN OF TRANSFORMED VALUES	*	MAR00530
*	SSE - SUM OF SQUARED DIFFERENCES, ERROR	*	MAR00540
*	(X AND X-BAR)	*	MAR00550
*	SST - SUM OF SQUARED DIFFERENCES, TOTAL	*	MAR00560
*	(X AND XBB)	*	MAR00570
*	SSB - SUM OF SQUARED DIFFERENCES, BETWEEN	*	MAR00580
*	(SST-SSE)	*	MAR00590
*	SHR - JAMES-STEIN SHRINKAGE VALUES	*	MAR00600
*	SHR1 - MODIFIED JAMES-STEIN SHRINKAGE	*	MAR00610
*	VALUES	*	MAR00620
*	SHREB - EMPIRICAL BAYES SHRINKAGE VALUES	*	MAR00630
*	SHRTEB - MODIFIED EMPIRICAL BAYES SHRINKAGE	*	MAR00640
*	VALUES	*	MAR00650
*	THE FOLLOWING ARE MODIFIED BASIC VARIABLES IN THE	*	MAR00660
*	TRANSFORMED SPACE. THIS IS DONE BY PLACING AN 'X' FOR	*	MAR00670
*	A PURE TRANSFORMED VARIABLE, OR AN 'R' FOR A RISK	*	MAR00680
*	VARIABLE IN FRONT OF THE BASIC VARIABLE.	*	MAR00690
*	MLE RML (MLE)	*	MAR00700
*	TSCA RTSCA	*	MAR00710
*	XJ RJ	*	MAR00720
*	XJ1 RJ1	*	MAR00730
*	XJLT RJLT	*	MAR00740
*	XEB RXEB	*	MAR00750
*	XTEB RTEB	*	MAR00760
*	ORIGINAL SPACE VARIABLES	*	MAR00770
*	THE FOLLOWING ARE MODIFIED BASIC VARIABLES IN THE	*	MAR00780
*	ORIGINAL SPACE. SIMILAR TO WHAT IS DONE IN THE	*	MAR00790
*	TRANSFORMED SPACE, A 'P' SIGNIFIES A PURE INVERTED	*	MAR00800
*	VARIABLE, AND AN 'RR' REPRESENTS A RISK VARIABLE. AN	*	MAR00810
*	'M' AT THE END OF A VARIABLE NAME IDENTIFIES A MILLER	*	MAR00820
*	INVERSE VARIABLE.	*	MAR00830
*	PMLE RRML (MLE)	*	MAR00840
*	PJ RRJ	*	MAR00850
*	PJM RRPJM	*	MAR00860
*	PJ1 RRJ1	*	MAR00870
*	PJ1M RRJ1M	*	MAR00880
*	PJLT RRJLT	*	MAR00890
*	PEB RRXEB	*	MAR00900
*	PEBM RRXEBM	*	MAR00910
*	PTEB RRRTEB	*	MAR00920
*	PTEBM RRTTEBM	*	MAR00930
*	THE FOLLOWING LIST IS OF PREFIXES USED TO DESCRIBE	*	MAR00940
*	THE MEAN ABSOLUTE DEVIATION VARIABLES:	*	MAR00950
*	MO - MEAN DEVIATION 'OVERAGE'	*	MAR00960
*	MU - MEAN DEVIATION 'UNDERAGE'	*	MAR00970
*	MD - MEAN 'TOTAL DEVIATION	*	MAR00980
*	KVO - NUMBER OF 'OVERAGE' ESTIMATES IN THE	*	MAR00990
*	AGGREGATE	*	MAR01000
*	KVU - NUMBER OF 'UNDERAGE' ESTIMATES IN THE	*	MAR01010
*	AGGREGATE	*	MAR01020
*	KVD - NUMBER OF 'TOTAL' ESTIMATION DEVIATIONS	*	MAR01030
*	IN THE AGGREGATE	*	MAR01040
*	THESE PREFIXES ARE APPENDED TO THE BASIC VARIABLES TO	*	MAR01050

```

*          FORM THE MAD VARIABLES AS FOLLOWS
*          MOTSCA MUTSCA MDTSCA KVOTSA KVUTSA KVD TSA (TSCA)
*          MOTSAM MUTSAM MDTSAM KVOTSM KVUTSM KVD TSM (TSCAM)
*          MOPMLE MUPMLE MDPMLE KVOPML KVUPML KVD PML (PMLE)
*          MOPJ MUPJ MDPJ KVOPJ KVUPJ KVD PJ (PJ)
*          MOPJM MUPJM MDPJM KVOPJM KVUPJM KVD PJM (PJM)
*          MOPJ1 MUPJ1 MDPJ1 KVOPJ1 KVUPJ1 KVD PJ1 (PJ1)
*          MOPJ1M MUPJ1M MDPJ1M KVOP1M KVUP1M KVD PJ1M (PJ1M)
*          MOPEB MUPEB MDPEB KVOPEB KVUPEB KVD PEB (PEB)
*          MOPEBM MUPEBM MDPEBM KVOPBM KVUPBM KVD PBM (PEBM)
*          MOPTB MUPTB MDPTB KVOPTB KVUPTB KVD PTB (PTB)
*          MOPTBM MUPTBM MDPTBM KVOBTM KVUBTM KVDBTM (PTEBM)
*
* INPUT/OUTPUT:
* INPUT IS TAKEN FROM A FILE CREATED FROM A TAPE
* PROVIDED BY NPRDC
* OUTPUT IS TO THE PRINTER
*
*****
**      FIXED PARAMETERS
PARAMETER (MXYR=7, MXLOS=31, MXGRD=10, MXMOS=15)
**      INPUT PARAMETERS
INTEGER ST1, ST2, LYR
INTEGER SLOS1, SLOS2
INTEGER SMOS(MXMOS), NMOS
INTEGER SGRD(MXGRD), NGRD
**      VARIABLES IN ORIGINAL SPACE
INTEGER T, FLAG, IVYR(MXYR)
INTEGER*2 D(MXMOS, MXLOS, MXGRD)
INTEGER KV(MXYR)
REAL AVCINV, AINV, ATT, ATT1
REAL CINV(MXMOS, MXLOS, MXGRD, MXYR), Y(MXMOS, MXLOS, MXGRD, MXYR)
REAL ACINV(MXMOS, MXLOS, MXGRD)
REAL PJ(MXMOS, MXLOS, MXGRD), PJM(MXMOS, MXLOS, MXGRD)
REAL PMLE(MXMOS, MXLOS, MXGRD), PJLT(MXMOS, MXLOS, MXGRD)
REAL PJ1(MXMOS, MXLOS, MXGRD), PJ1M(MXMOS, MXLOS, MXGRD)
REAL PEB(MXMOS, MXLOS, MXGRD), PEBM(MXMOS, MXLOS, MXGRD)
REAL PTEB(MXMOS, MXLOS, MXGRD), PTEBM(MXMOS, MXLOS, MXGRD)
REAL RRTSCA(MXYR), RRJ(MXYR), RRJLT(MXYR), RRML(MXYR)
REAL RRTSAM(MXYR), RRJM(MXYR), RRJ1(MXYR), RRJ1M(MXYR)
REAL RRPEB(MXYR), RRPEBM(MXYR)
REAL RRPTEB(MXYR), RRPTEBM(MXYR)
**      VARIABLES IN TRANSFORMED SPACE
REAL X(MXMOS, MXLOS, MXGRD, MXYR)
REAL XT(MXMOS, MXLOS, MXGRD, MXYR)
REAL XB(MXMOS, MXLOS, MXGRD), TSCA(MXMOS, MXLOS, MXGRD)
REAL XTB(MXMOS, MXLOS, MXGRD)
REAL XJ(MXMOS, MXLOS, MXGRD), XJLT(MXMOS, MXLOS, MXGRD)
REAL XEB(MXMOS, MXLOS, MXGRD), XJ1(MXMOS, MXLOS, MXGRD)
REAL XTEB(MXMOS, MXLOS, MXGRD)
REAL XMLE(MXMOS, MXLOS, MXGRD)
REAL SHREB(MXMOS, MXLOS, MXGRD)
REAL SHRTEB(MXMOS, MXLOS, MXGRD)
REAL RML(MXYR), RSL(MXYR), RXEB(MXYR)
REAL RXTEB(MXYR)
REAL RTSCA(MXYR), RTJ(MXYR), RTJLT(MXYR)
REAL RJ(MXYR), RJLT(MXYR), RJ1(MXYR)
REAL TSCAM(MXMOS, MXLOS, MXGRD)

```

REAL V(MXMOS,MXLOS,MXGRD)	MAR01620
REAL VT(MXMOS,MXLOS,MXGRD)	MAR01630
REAL*8 SHR, SHRX, SHR1, SSB, SST, SSE, XBB	MAR01640
** VARIABLES FOR FIGURE OF MERIT	MAR01650
INTEGER KVOTSA(MXYR),KVOTSM(MXYR),KVOPJ(MXYR),KVOPJ1(MXYR)	MAR01660
INTEGER KVOPJM(MXYR),KVOPEB(MXYR),KVOP1M(MXYR),KVOPBM(MXYR)	MAR01670
INTEGER KVOPTB(MXYR),KVOBTM(MXYR)	MAR01680
INTEGER KVOPML(MXYR)	MAR01690
INTEGER KVUTSA(MXYR),KVUTSM(MXYR),KVUPJ(MXYR),KVUPJ1(MXYR)	MAR01700
INTEGER KVUPJM(MXYR),KVUPEB(MXYR),KVUP1M(MXYR),KVUPBM(MXYR)	MAR01710
INTEGER KVUPTB(MXYR),KVUBTM(MXYR)	MAR01720
INTEGER KVUPML(MXYR)	MAR01730
INTEGER KVD TSA(MXYR),KVD TSM(MXYR),KVDPJ(MXYR),KVDPJ1(MXYR)	MAR01740
INTEGER KVDPJM(MXYR),KVDP EB(MXYR),KVDP 1M(MXYR),KVDPBM(MXYR)	MAR01750
INTEGER KVDPTB(MXYR),KVDBTM(MXYR)	MAR01760
INTEGER KVDPML(MXYR)	MAR01770
REAL MUTSCA(MXYR),MUTSAM(MXYR),MUPJ(MXYR),MUPJ1(MXYR),MUPJM(MXYR)	MAR01780
REAL MUPJ1M(MXYR),MUPEBM(MXYR),MUPEB(MXYR),MUPMLE(MXYR)	MAR01790
REAL MUPTBM(MXYR),MUPT EB(MXYR)	MAR01800
REAL MOTSCA(MXYR),MOTSAM(MXYR),MOPJ(MXYR),MOPJ1(MXYR),MOPJM(MXYR)	MAR01810
REAL MOPJ1M(MXYR),MOPEBM(MXYR),MOPEB(MXYR),MOPMLE(MXYR)	MAR01820
REAL MOPTBM(MXYR),MOPT EB(MXYR)	MAR01830
REAL MDTSCA(MXYR),MDTSAM(MXYR),MDPJ(MXYR),MDPJ1(MXYR),MDPJM(MXYR)	MAR01840
REAL MDPJ1M(MXYR),MDPEBM(MXYR),MDPEB(MXYR),MDPMLE(MXYR)	MAR01850
REAL MDPTBM(MXYR),MDPT EB(MXYR)	MAR01860
REAL MILLER	MAR01870
** INPUT DATA AREAS	MAR01880
INTEGER MOS, LOS, GRADE, YR, INV	MAR01890
INTEGER MOS1,LOS1,GRADE1,YR1,INV1	MAR01900
** INPUT INITIALIZATION	MAR01910
DATA KV/MXYR*0/	MAR01920
DATA MOPMLE/MXYR*0/,MUPMLE/MXYR*0/,MDPML/MXYR*0/	MAR01930
DATA KVOPML/MXYR*0/,KVUPML/MXYR*0/,KVDPML/MXYR*0/	MAR01940
DATA MOTSCA/MXYR*0/,MUTSCA/MXYR*0/,MDTSCA/MXYR*0/	MAR01950
DATA KVOTSA/MXYR*0/,KVUTSA/MXYR*0/,KVD TSA/MXYR*0/	MAR01960
DATA MOTSAM/MXYR*0/,MUTSAM/MXYR*0/,MDTSAM/MXYR*0/	MAR01970
DATA KVOTSM/MXYR*0/,KVUTSM/MXYR*0/,KVD TSM/MXYR*0/	MAR01980
DATA MOPJ/MXYR*0/,MUPJ/MXYR*0/,MDPJ/MXYR*0/	MAR01990
DATA KVOPJ/MXYR*0/,KVUPJ/MXYR*0/,KVDPJ/MXYR*0/	MAR02000
DATA MOPJM/MXYR*0/,MUPJM/MXYR*0/,MDPJM/MXYR*0/	MAR02010
DATA KVOPJM/MXYR*0/,KVUPJM/MXYR*0/,KVDPJM/MXYR*0/	MAR02020
DATA MOPJ1/MXYR*0/,MUPJ1/MXYR*0/,MDPJ1/MXYR*0/	MAR02030
DATA KVOPJ1/MXYR*0/,KVUPJ1/MXYR*0/,KVDPJ1/MXYR*0/	MAR02040
DATA MOPJ1M/MXYR*0/,MUPJ1M/MXYR*0/,MDPJ1M/MXYR*0/	MAR02050
DATA KVOP1M/MXYR*0/,KVUP1M/MXYR*0/,KVDP1M/MXYR*0/	MAR02060
DATA KVOPEB/MXYR*0/,MUPEB/MXYR*0/,MDPEB/MXYR*0/	MAR02070
DATA KVOPEBM/MXYR*0/,MUPEBM/MXYR*0/,MDPEBM/MXYR*0/	MAR02080
DATA KVOPEBM/MXYR*0/,MUPEBM/MXYR*0/,MDPEBM/MXYR*0/	MAR02090
DATA KVOPBM/MXYR*0/,KVUPBM/MXYR*0/,KVDPBM/MXYR*0/	MAR02100
DATA MOPT EB/MXYR*0/,MUPT EB/MXYR*0/,MDPT EB/MXYR*0/	MAR02110
DATA KVOPTB/MXYR*0/,KVUPTB/MXYR*0/,KVDPTB/MXYR*0/	MAR02120
DATA MOPTBM/MXYR*0/,MUPTBM/MXYR*0/,MDPTBM/MXYR*0/	MAR02130
DATA KVOBTM/MXYR*0/,KVUBTM/MXYR*0/,KVDBTM/MXYR*0/	MAR02140
*	MAR02150
*****	MAR02160

**	READ PARAMETER DATA TO USE IN THIS RUN.	**	MAR02170
*****		*****	MAR02180
	WRITE(5,100)		MAR02190
100	FORMAT(' ENTER 1ST YEAR AND LAST YEAR TO USE FOR ESTIMATION')		MAR02200
	READ(5,*) ST1,ST2		MAR02210
	WRITE(6,*) 'ESTIMATION YEARS: ',ST1,ST2		MAR02220
*			MAR02230
	WRITE(5,101)		MAR02240
101	FORMAT(' ENTER LAST YEAR PRESENT IN THE DATA BASE')		MAR02250
	READ(5,*) LYR		MAR02260
	WRITE(6,*) 'LAST YEAR AVAILABLE FOR VALIDATION: ',LYR		MAR02270
*			MAR02280
	WRITE(5,102)		MAR02290
102	FORMAT(' ENTER NO. OF MOS FOLLOWED BY ARRAY OF SAME LENGTH')		MAR02300
	READ(5,*) NMOS, (SMOS(I), I=1,NMOS)		MAR02310
	WRITE(6,*) 'MOS SELECTED: ', (SMOS(I), I=1,NMOS)		MAR02320
*			MAR02330
	WRITE(5,103)		MAR02340
103	FORMAT(' ENTER 1ST AND LAST LOS VALUE TO USE')		MAR02350
	READ(5,*) SLOS1,SLOS2		MAR02360
	WRITE(6,*) 'LOS RANGE: ',SLOS1,SLOS2		MAR02370
*			MAR02380
	WRITE(5,104)		MAR02390
104	FORMAT(' ENTER NO. OF GRADE FOLLOWED BY ARRAY OF SAME LENGTH')		MAR02400
	READ(5,*) NGRD, (SGRD(I), I=1,NGRD)		MAR02410
	WRITE(6,*) 'GRADES SELECTED', (SGRD(I), I=1,NGRD)		MAR02420
*			MAR02430
	WRITE(5,105)		MAR02440
105	FORMAT(' ENTER DEE FACTOR')		MAR02450
	READ(5,*) DEE		MAR02460
	WRITE(6,*) 'DEE FACTOR: ',DEE		MAR02470
*			MAR02480
**	COMPUTE START AND END YEARS FOR VALIDATION		MAR02490
	NYR=LYR-ST1+1		MAR02500
	NVYR=LYR-ST2		MAR02510
	NLOS=SLOS2-SLOS1+1		MAR02520
	T=ST2-ST1+1		MAR02530
	DO 41 I=1,NVYR		MAR02540
	IVYR(I)=ST2+I		MAR02550
41	CONTINUE		MAR02560
*****		*****	MAR02570
**	INITIALIZE D ARRAY TO ALLOW FINDING MISSING VALUES.	**	MAR02580
*****		*****	MAR02590
	DO 1 I=1,MXMOS		MAR02600
	DO 2 J=1,MXLOS		MAR02610
	DO 3 K=1,MXGRD		MAR02620
	D(I,J,K)=-9999		MAR02630
3	CONTINUE		MAR02640
2	CONTINUE		MAR02650
1	CONTINUE		MAR02660
*****		*****	MAR02670
**	TAPE PROCESSING TO INPUT INVENTORY AND ATTRITION VALUES,	**	MAR02680
**	DEVELOPE D MATRIX ACCORDING TO EXTRACTION CRITERIA.	**	MAR02690
*****		*****	MAR02700
	FLAG=0		MAR02710
	IEOF=0		MAR02720

	CALL READER(MOS1,LOS1,GRADE1,YR1,INV1,ATT1, NMOS,SMOS,	MAR02730
*	NGRD,SGRD, SLOS1,SLOS2, ST1,ST2, IM1,IL1,IG1,IT1,IEOF)	MAR02740
*	IF(IEOF .NE. 0) THEN	MAR02750
	WRITE(6,*) '***** NO DATA MEETS SELECTIONS REQS'	MAR02760
	STOP	MAR02770
	ENDIF	MAR02780
*		MAR02790
11	IF(IEOF .NE. 0) GO TO 9	MAR02800
	MOS=MOS1	MAR02810
	LOS=LOS1	MAR02820
	GRADE=GRADE1	MAR02830
	YR=YR1	MAR02840
	INV=INV1	MAR02850
	ATT=ATT1	MAR02860
	IM=IM1	MAR02870
	IL=IL1	MAR02880
	IG=IG1	MAR02890
	IT=IT1	MAR02900
	IF(YR.GT.ST1) FLAG = 1	MAR02910
**	CHECK CASE WHERE DATA BEGINS PAST 1ST YR	MAR02920
	IF(YR .GT. ST1) CINV(IM,IL,IG,IT-1)=.5*FLOAT(INV)	MAR02930
	CALL READER(MOS1,LOS1,GRADE1,YR1,INV1,ATT1, NMOS,SMOS,	MAR02940
*	NGRD,SGRD, SLOS1,SLOS2, ST1,ST2, IM1,IL1,IG1,IT1,IEOF)	MAR02950
**	ONLY FOR CHECKING PURPOSES. MARK D TO INDICATE SOME YR PRESENT	MAR02960
	D(IM,IL,IG)=0	MAR02970
*		MAR02980
12	IF(.NOT.(IEOF.EQ.0 .AND. MOS1.EQ.MOS .AND. LOS1.EQ.LOS	MAR02990
*	.AND. GRADE1.EQ.GRADE)) GO TO 8	MAR03000
**	CENTRAL INV. FOR YR1 - 1	MAR03010
	IF(YR1 .GT. YR+1) THEN	MAR03020
	CINV(IM1,IL1,IG1,IT1-1)=.5*FLOAT(INV1)	MAR03030
	AINV=.5*FLOAT(INV)	MAR03040
	I=YR+1	MAR03050
	FLAG=1	MAR03060
	ELSE	MAR03070
	AINV=.5*FLOAT(INV+INV1)	MAR03080
	END IF	MAR03090
	CINV(IM,IL,IG,IT)=AMAX1(AINV,ATT)	MAR03100
	Y(IM,IL,IG,IT)=ATT	MAR03110
	YR=YR1	MAR03120
	INV=INV1	MAR03130
	ATT=ATT1	MAR03140
	IM=IM1	MAR03150
	IL=IL1	MAR03160
	IG=IG1	MAR03170
	IT=IT1	MAR03180
	CALL READER(MOS1,LOS1,GRADE1,YR1,INV1,ATT1, NMOS,SMOS,	MAR03190
*	NGRD,SGRD, SLOS1,SLOS2, ST1,ST2, IM1,IL1,IG1,IT1,IEOF)	MAR03200
	GO TO 12	MAR03210
8	CONTINUE	MAR03220
**	WHEN YEARS MISSING AT THE END	MAR03230
	IF(YR .LT. LYR) THEN	MAR03240
	FLAG=1	MAR03250
	AINV=.5*FLOAT(INV)	MAR03260
	ELSE	MAR03270
		MAR03280

AINV=FLOAT(INV)	MAR03290
END IF	MAR03300
CINV(IM,IL,IG,IT)=AMAX1(AINV,ATT)	MAR03310
Y(IM,IL,IG,IT)=ATT	MAR03320
*	MAR03330
CC=CINV(IM,IL,IG,IT)	MAR03340
TEMP=-1.+2.*ATT/(1.+CC)	MAR03350
TEMP1=-1.+2.*(1.+ATT)/(1.+CC)	MAR03360
CC=ABS(TEMP)	MAR03370
CC1=ABS(TEMP1)	MAR03380
IF(CC.GT.1..OR.CC1.GT.1.) WRITE(6,*) '*** TEMP, TEMP1='	MAR03390
* TEMP,TEMP1,'***MOS,LOS,GR,YR=',SMOS(IM),IL,SGRD(IG),IT	MAR03400
*	MAR03410
GO TO 11	MAR03420
9 CONTINUE	MAR03430
*****	MAR03440
** CALCULATE AVERAGE CENTRAL INVENTORY AND MAXIMUM	MAR03450
** LIKELIHOOD ESTIMATE.	MAR03460
*****	MAR03470
DO 10 IM=1,NMOS	MAR03480
DO 20 IL=1,NLOS	MAR03490
DO 30 IG=1,NGRD	MAR03500
SUMY=0	MAR03510
ACINV(IM,IL,IG)=0	MAR03520
DO 35 IT=1,T	MAR03530
SUMY=SUMY+Y(IM,IL,IG,IT)	MAR03540
ACINV(IM,IL,IG)=ACINV(IM,IL,IG)+CINV(IM,IL,IG,IT)	MAR03550
35 CONTINUE	MAR03560
IF(ACINV(IM,IL,IG).NE.0) THEN	MAR03570
PMLE(IM,IL,IG)=SUMY/ACINV(IM,IL,IG)	MAR03580
ELSE	MAR03590
PMLE(IM,IL,IG)=0	MAR03600
ENDIF	MAR03610
ACINV(IM,IL,IG)=ACINV(IM,IL,IG)/T	MAR03620
DO 40 IT=1,NYR	MAR03630
** IF ANY CENTRAL INVENTORY>0 (OVER EST. YRS T) THEN D=1	MAR03640
IF(CINV(IM,IL,IG,IT).GT.0.AND.IT.LE.T) D(IM,IL,IG)=1	MAR03650
** IF ANY CENTRAL INVENTORY>0 (OVER VAL. YRS > T) THEN COUNT	MAR03660
IF(CINV(IM,IL,IG,IT).GT.0.AND.IT.GT.T) KV(IT-T)=KV(IT-T)+1	MAR03670
** TRANSFORMATION OF CINV USING THE FREEMAN-TUKEY	MAR03680
** DOUBLE ARCSIN FORMULA.	MAR03690
CALL FTT(CINV(IM,IL,IG,IT),Y(IM,IL,IG,IT),X(IM,IL,IG,IT))	MAR03700
XT(IM,IL,IG,IT)=X(IM,IL,IG,IT) / SQRT(0.5+CINV(IM,IL,IG,IT))	MAR03710
40 CONTINUE	MAR03720
30 CONTINUE	MAR03730
20 CONTINUE	MAR03740
10 CONTINUE	MAR03750
*****	MAR03760
** CHECK FOR MISSING DATA, COMPUTE THE TRANSFORMED	MAR03770
** CELL AVERAGES AND GRAND MEAN.	MAR03780
*****	MAR03790
XBB=0	MAR03800
KK=0	MAR03810
DO 50 IM=1,NMOS	MAR03820
DO 60 IL=1,NLOS	MAR03830
DO 70 IG=1,NGRD	MAR03840

**	IF(D(IM,IL,IG) .EQ. -9999) THEN	MAR03850
	REPORT MISSING COMBINATION IM,IL,IG & CLEAR D ENTRY	MAR03860
	D(IM,IL,IG)=0	MAR03870
	ILD=SLOS1+IL-1	MAR03880
	FLAG=1	MAR03890
	END IF	MAR03900
	KK=KK+D(IM,IL,IG)	MAR03910
	XB(IM,IL,IG)=0	MAR03920
	XTB(IM,IL,IG)=0	MAR03930
	DO 80 IT=1,T	MAR03940
	XB(IM,IL,IG)=XB(IM,IL,IG)+X(IM,IL,IG,IT)	MAR03950
	XTB(IM,IL,IG)=XTB(IM,IL,IG)+XT(IM,IL,IG,IT)	MAR03960
80	CONTINUE	MAR03970
	XB(IM,IL,IG)=XB(IM,IL,IG)/T	MAR03980
	XTB(IM,IL,IG)=XTB(IM,IL,IG)/T	MAR03990
**	EMPIRICAL BAYES PREPARATION ON THE TRANSFORMED SCALE.	MAR04000
	XEB(IM,IL,IG)=XB(IM,IL,IG)	MAR04010
	XTEB(IM,IL,IG)=XTB(IM,IL,IG)	MAR04020
*		MAR04030
	XBB=XBB+D(IM,IL,IG)*XB(IM,IL,IG)	MAR04040
70	CONTINUE	MAR04050
60	CONTINUE	MAR04060
50	CONTINUE	MAR04070
	XBB=XBB/KK	MAR04080
*****	*****	MAR04090
**	BAYES VARIANCE ITERATION SECTION.	MAR04100
*****	*****	MAR04110
	IC=0	MAR04120
**	COMPUTE TRANSFORMED SCALE CELL VARIANCE.	MAR04130
499	DO 500 IM=1,NMOS	MAR04140
	DO 510 IL=1,NLOS	MAR04150
	DO 520 IG=1,NGRD	MAR04160
	IF (D(IM,IL,IG).NE.1) GOTO 520	MAR04170
	AVSQR = 0.0	MAR04180
	DO 522 IT=1,T	MAR04190
522	AVSQR = AVSQR + SQRT(0.5 + CINV(IM,IL,IG,IT))	MAR04200
	Z = XEB(IM,IL,IG) + (AVSQR/T) * 1.5708	MAR04210
	V(IM,IL,IG) = VAR(Z)	MAR04220
520	CONTINUE	MAR04230
510	CONTINUE	MAR04240
500	CONTINUE	MAR04250
*		MAR04260
	A=0	MAR04270
530	AI=A	MAR04280
	SALPH=0	MAR04290
	DO 531 IM=1,NMOS	MAR04300
	DO 532 IL=1,NLOS	MAR04310
	DO 533 IG=1,NGRD	MAR04320
533	IF (D(IM,IL,IG).EQ.1) SALPH = SALPH + 1./(A + V(IM,IL,IG))	MAR04330
532	CONTINUE	MAR04340
531	CONTINUE	MAR04350
*		MAR04360
	ZB=0	MAR04370
**	COMPUTE BAYES ALPHA AND GAMMA PARAMETERS	MAR04380
	DO 537 IM=1,NMOS	MAR04390
	DO 538 IL=1,NLOS	MAR04400

	DO 539 IG=1,NGRD	MAR04410
	IF (D(IM,IL,IG).EQ.1) THEN	MAR04420
	ALPH = 1./(A + V(IM,IL,IG))	MAR04430
	GAM = ALPH / SALPH	MAR04440
	ZB = ZB + GAM*XB(IM,IL,IG)	MAR04450
	ENDIF	MAR04460
539	CONTINUE	MAR04470
538	CONTINUE	MAR04480
537	CONTINUE	MAR04490
	F=0	MAR04500
	G=0	MAR04510
	DO 540 IM=1,NMOS	MAR04520
	DO 541 IL=1,NLOS	MAR04530
	DO 542 IG=1,NGRD	MAR04540
	IF (D(IM,IL,IG).EQ.1) THEN	MAR04550
	ALPH = 1./(A + V(IM,IL,IG))	MAR04560
	F = F + (ALPH * (XB(IM,IL,IG) - ZB)**2)	MAR04570
	G = G + (ALPH * (XB(IM,IL,IG) - ZB)**2)	MAR04580
	ENDIF	MAR04590
542	CONTINUE	MAR04600
541	CONTINUE	MAR04610
540	CONTINUE	MAR04620
**	COMPUTE THE PRIOR VARIANCE (A)	MAR04630
	A = A - (KK-1-F)/G	MAR04640
*		MAR04650
	IF (A.LE.0.) THEN	MAR04660
	A=0	MAR04670
	GOTO 550	MAR04680
	ENDIF	MAR04690
	IF (ABS(A-AI).GT..0001) GOTO 530	MAR04700
**	COMPUTE THE EMPIRICAL BAYES ATTRITION RATES	MAR04710
550	DO 570 IM=1,NMOS	MAR04720
	DO 580 IL=1,NLOS	MAR04730
	DO 590 IG=1,NGRD	MAR04740
	IF (D(IM,IL,IG).EQ.1) THEN	MAR04750
	DEN = A + V(IM,IL,IG)	MAR04760
	XEB(IM,IL,IG) = (A/DEN)*XB(IM,IL,IG)+((DEN-A)/DEN)*ZB	MAR04770
	ENDIF	MAR04780
590	CONTINUE	MAR04790
580	CONTINUE	MAR04800
570	CONTINUE	MAR04810
	IC=IC+1	MAR04820
	IF (A.EQ.0) GOTO 599	MAR04830
	IF (IC.LT.10) GOTO 499	MAR04840
*		MAR04850
599	CONTINUE	MAR04860
*		MAR04870
**	COMPUTE THE EMPIRICAL BAYES SHRINKAGE RATES	MAR04880
	DO 600 IM=1,NMOS	MAR04890
	DO 601 IL=1,NLOS	MAR04900
	DO 602 IG=1,NGRD	MAR04910
	SHREB(IM,IL,IG) = 0	MAR04920
	IF (D(IM,IL,IG).EQ.1) THEN	MAR04930
	SHREB(IM,IL,IG) = V(IM,IL,IG)/(A + V(IM,IL,IG))	MAR04940
	ENDIF	MAR04950
602	CONTINUE	MAR04960

601	CONTINUE	MAR04970
600	CONTINUE	MAR04980
*****		MAR04990
**	ALTERNATE BAYES VARIANCE ITERATION SECTION.	MAR05000
*****		MAR05010
**	COMPUTE TRANSFORMED SCALE CELL VARIANCE.	MAR05020
699	DO 700 IM=1,NMOS	MAR05030
	DO 710 IL=1,NLOS	MAR05040
	DO 720 IG=1,NGRD	MAR05050
	VT(IM,IL,IG) = 0.0	MAR05060
	IF (D(IM,IL,IG).NE.1) GOTO 720	MAR05070
	DO 725 IT=1,T	MAR05080
	ZT = X(IM,IL,IG,IT) + 1.5708 * SQRT(0.5 +	MAR05090
*	CINV(IM,IL,IG,IT))	MAR05100
725	VT(IM,IL,IG) = VT(IM,IL,IG) + VAR(ZT)/(0.5 +	MAR05110
*	CINV(IM,IL,IG,IT))	MAR05120
	VT(IM,IL,IG) = VT(IM,IL,IG)/T**2	MAR05130
720	CONTINUE	MAR05140
710	CONTINUE	MAR05150
700	CONTINUE	MAR05160
*		MAR05170
	AT=0	MAR05180
730	ATI=AT	MAR05190
	SALPHT=0	MAR05200
	DO 731 IM=1,NMOS	MAR05210
	DO 732 IL=1,NLOS	MAR05220
	DO 733 IG=1,NGRD	MAR05230
733	IF (D(IM,IL,IG).EQ.1) SALPHT= SALPHT+ 1./(AT+ VT(IM,IL,IG))	MAR05240
732	CONTINUE	MAR05250
731	CONTINUE	MAR05260
	ZTB=0	MAR05270
**	COMPUTE BAYES ALPHA AND GAMMA PARAMETERS	MAR05280
	DO 737 IM=1,NMOS	MAR05290
	DO 738 IL=1,NLOS	MAR05300
	DO 739 IG=1,NGRD	MAR05310
	IF (D(IM,IL,IG).EQ.1) THEN	MAR05320
	ALPHT= 1./(AT+ VT(IM,IL,IG))	MAR05330
	GAMT = ALPHT / SALPHT	MAR05340
	ZTB = ZTB+ GAMT*XTB(IM,IL,IG)	MAR05350
	ENDIF	MAR05360
739	CONTINUE	MAR05370
738	CONTINUE	MAR05380
737	CONTINUE	MAR05390
	PZTB=0.5*(1.+SIN(ZTB))	MAR05400
	FT=0	MAR05410
	GT=0	MAR05420
	DO 740 IM=1,NMOS	MAR05430
	DO 741 IL=1,NLOS	MAR05440
	DO 742 IG=1,NGRD	MAR05450
	IF (D(IM,IL,IG).EQ.1) THEN	MAR05460
	ALPHT= 1./(AT + VT(IM,IL,IG))	MAR05470
	FT = FT + (ALPHT * (XTB(IM,IL,IG) - ZTB)**2)	MAR05480
	GT = GT + (ALPHT * (XTB(IM,IL,IG) - ZTB)**2)	MAR05490
	ENDIF	MAR05500
742	CONTINUE	MAR05510
741	CONTINUE	MAR05520

740	CONTINUE	MAR05530
**	COMPUTE THE PRIOR VARIANCE (A)	MAR05540
	AT= AT- (KK-1-FT)/GT	MAR05550
*		MAR05560
	IF (AT.LE.O.) THEN	MAR05570
	AT=0	MAR05580
	GOTO 750	MAR05590
	ENDIF	MAR05600
	IF (ABS(AT-ATI).GT..0001) GOTO 730	MAR05610
**	COMPUTE THE EMPIRICAL BAYES ATTRITION RATES	MAR05620
750	DO 770 IM=1,NMOS	MAR05630
	DO 780 IL=1,NLOS	MAR05640
	DO 790 IG=1,NGRD	MAR05650
	IF (D(IM,IL,IG).EQ.1) THEN	MAR05660
	DENT = AT + VT(IM,IL,IG)	MAR05670
	XTEB(IM,IL,IG) = (AT/DENT)*XTB(IM,IL,IG) +	MAR05680
*	((DENT-AT)/DENT)*ZTB	MAR05690
	ENDIF	MAR05700
790	CONTINUE	MAR05710
780	CONTINUE	MAR05720
770	CONTINUE	MAR05730
**	COMPUTE THE EMPIRICAL BAYES SHRINKAGE RATES	MAR05740
	DO 800 IM=1,NMOS	MAR05750
	DO 801 IL=1,NLOS	MAR05760
	DO 802 IG=1,NGRD	MAR05770
	SHRTEB(IM,IL,IG) = 0	MAR05780
	IF (D(IM,IL,IG).EQ.1) THEN	MAR05790
	SHRTEB(IM,IL,IG) = VT(IM,IL,IG)/(AT + VT(IM,IL,IG))	MAR05800
	ENDIF	MAR05810
802	CONTINUE	MAR05820
801	CONTINUE	MAR05830
800	CONTINUE	MAR05840
*****		MAR05850
**	JAMES-STEIN SHRINKAGE RATES.	MAR05860
*****		MAR05870
	SST=0	MAR05880
	SSE=0	MAR05890
	SSB=0	MAR05900
**	COMPUTE THE SUMS OF SQUARES (SST, SSE AND SSB)	MAR05910
	DO 90 IM=1,NMOS	MAR05920
	DO 95 IL=1,NLOS	MAR05930
	DO 110 IG=1,NGRD	MAR05940
	TEMP1=0	MAR05950
	TEMP2=0	MAR05960
	DO 120 IT=1,T	MAR05970
	TEMP1=TEMP1+(X(IM,IL,IG,IT)-XBB)**2	MAR05980
	TEMP2=TEMP2+(X(IM,IL,IG,IT)-XB(IM,IL,IG))**2	MAR05990
120	CONTINUE	MAR06000
	SST=SST+D(IM,IL,IG)*TEMP1	MAR06010
	SSE=SSE+D(IM,IL,IG)*TEMP2	MAR06020
110	CONTINUE	MAR06030
95	CONTINUE	MAR06040
90	CONTINUE	MAR06050
	SSB=SST-SSE	MAR06060
**	COMPUTE THE JAMES-STEIN SHRINKAGE RATES	MAR06070
	SHRX=(SSE/SSB)* (KK-3)/(KK*(T-1)+2)	MAR06080

SHR=DMIN1(SHRX, 1.DO)	MAR06090
SHR=DMAX1(SHR, 0.DO)	MAR06100
SHR1=T*SHR	MAR06110
*****	MAR06120
** LIMITED TRANSLATION OPTION & JAMES-STEIN **	MAR06130
** ON TRANSFORMED SCALE. **	MAR06140
*****	MAR06150
NN=KK*(T-1)	MAR06160
DO 125 IM=1,NMOS	MAR06170
DO 130 IL=1,NLOS	MAR06180
DO 140 IG=1,NGRD	MAR06190
XJ(IM,IL,IG)=0	MAR06200
XJ1(IM,IL,IG)=0	MAR06210
IF (D(IM,IL,IG).EQ.1) THEN	MAR06220
XJ(IM,IL,IG)=XBB+(1-SHR)*(XB(IM,IL,IG)-XBB)	MAR06230
XJ1(IM,IL,IG)=XBB+(1-SHR1)*(XB(IM,IL,IG)-XBB)	MAR06240
C *** XJLT(IM,IL,IG)=XBB+(1-SHR*RHO(XB(IM,IL,IG),XBB,SSE,	MAR06250
C *** SSB,DEE,KK,T,NN))*(XB(IM,IL,IG)-XBB)	MAR06260
XMLE(IM,IL,IG)=SQRT(.5+ACINV(IM,IL,IG))	MAR06270
* *ASIN(2*PMLE(IM,IL,IG)-1)	MAR06280
ENDIF	MAR06290
140 CONTINUE	MAR06300
130 CONTINUE	MAR06310
125 CONTINUE	MAR06320
*****	MAR06330
** VALIDATION. RISK EVALUATED IN TRANSFORM SPACE. **	MAR06340
*****	MAR06350
DO 150 IT=1, NVYR	MAR06360
RTSCA(IT)=0	MAR06370
RJ(IT)=0	MAR06380
RJ1(IT)=0	MAR06390
RXEB(IT)=0	MAR06400
RXTEB(IT)=0	MAR06410
C *** RJLT(IT)=0	MAR06420
RML(IT)=0	MAR06430
DO 160 IM=1,NMOS	MAR06440
DO 170 IL=1,NLOS	MAR06450
DO 180 IG=1,NGRD	MAR06460
IF(D(IM,IL,IG).EQ.1 .AND. CINV(IM,IL,IG,IT+T).NE.0) THEN	MAR06470
RTSCA(IT)=RTSCA(IT)+(X(IM,IL,IG,T+IT)- XB(IM,IL,IG))**2	MAR06480
RJ(IT)= RJ(IT)+(X(IM,IL,IG,T+IT)- XJ(IM,IL,IG))**2	MAR06490
RJ1(IT)= RJ1(IT)+(X(IM,IL,IG,T+IT)- XJ1(IM,IL,IG))**2	MAR06500
RXEB(IT)=RXEB(IT)+(X(IM,IL,IG,T+IT)- XEB(IM,IL,IG))**2	MAR06510
RXTEB(IT)=RXTEB(IT)+((X(IM,IL,IG,T+IT)/SQRT(0.5 +	MAR06520
* CINV(IM,IL,IG,IT+T))- XTEB(IM,IL,IG))**2)	MAR06530
* /VT(IM,IL,IG)	MAR06540
C *** RJLT (IT)= RJLT(IT)+(X(IM,IL,IG,T+IT)-XJLT(IM,IL,IG))**2	MAR06550
RML (IT)= RML (IT)+(X(IM,IL,IG,T+IT)-XMLE(IM,IL,IG))**2	MAR06560
ENDIF	MAR06570
180 CONTINUE	MAR06580
170 CONTINUE	MAR06590
160 CONTINUE	MAR06600
RTSCA(IT)=RTSCA(IT)/KV(IT)	MAR06610
RJ (IT)=RJ (IT)/KV(IT)	MAR06620
RJ1 (IT)=RJ1 (IT)/KV(IT)	MAR06630
RXEB(IT)=RXEB(IT)/KV(IT)	MAR06640

	RXTEB(IT)=RXTEB(IT)/(KV(IT)*T)	MAR06650
C ***	RJLT(IT)=RJLT(IT)/KV(IT)	MAR06660
	RML (IT)=RML (IT)/KV(IT)	MAR06670
	RTJ(IT)=RTSCA(IT)/RJ(IT)	MAR06680
C ***	RTJLT(IT)=RTSCA(IT)/RJLT(IT)	MAR06690
C ***	RSL(IT)=(RJ(IT)-RTSCA(IT))/(RJLT(IT)-RTSCA(IT))	MAR06700
150	CONTINUE	MAR06710
*****		MAR06720
**	DO RISK ANALYSIS ON THE ORIGINAL SCALE.	MAR06730
*****		MAR06740
	DO 200 IM=1,NMOS	MAR06750
	DO 210 IL=1,NLOS	MAR06760
	DO 220 IG=1,NGRD	MAR06770
**	AVERAGE CENTRAL INVENTORY OVER ESTIMATION SET	MAR06780
	CINVS=0	MAR06790
	DO 230 IT=1,T	MAR06800
	AVCINV=AVCINV + CINV(IM,IL,IG,IT)	MAR06810
230	CONTINUE	MAR06820
	AVCINV=AVCINV/T	MAR06830
	TSCA(IM,IL,IG)=D(IM,IL,IG)*SCINV(XB(IM,IL,IG), AVCINV)	MAR06840
	PJ(IM,IL,IG)=D(IM,IL,IG)*SCINV(XJ(IM,IL,IG), AVCINV)	MAR06850
	PJ1(IM,IL,IG)=D(IM,IL,IG)*SCINV(XJ1(IM,IL,IG), AVCINV)	MAR06860
	PEB(IM,IL,IG)=D(IM,IL,IG)*SCINV(XEB(IM,IL,IG), AVCINV)	MAR06870
	PTEB(IM,IL,IG)=D(IM,IL,IG)*0.5*(1. + SIN(XTEB(IM,IL,IG)))	MAR06880
C ***	PJLT(IM,IL,IG)=D(IM,IL,IG)*SCINV(XJLT(IM,IL,IG), AVCINV)	MAR06890
	IF (D(IM,IL,IG).EQ.1) THEN	MAR06900
	TSCAM(IM,IL,IG)=MILLER(XB(IM,IL,IG), AVCINV)	MAR06910
	PJM(IM,IL,IG)=MILLER(XJ(IM,IL,IG), AVCINV)	MAR06920
	PJ1M(IM,IL,IG)=MILLER(XJ1(IM,IL,IG), AVCINV)	MAR06930
	PEBM(IM,IL,IG)=MILLER(XEB(IM,IL,IG), AVCINV)	MAR06940
	XMIL = XTEB(IM,IL,IG) * SQRT(0.5 + AVCINV)	MAR06950
	PTEBM(IM,IL,IG)=MILLER(XMIL , AVCINV)	MAR06960
	ENDIF	MAR06970
220	CONTINUE	MAR06980
210	CONTINUE	MAR06990
200	CONTINUE	MAR07000
*****		MAR07010
**	VALIDATION. RISK EVALUATED IN ORIGINAL SPACE.	MAR07020
*****		MAR07030
	DO 240 IT=1, NVYR	MAR07040
	RRTSCA(IT)=0	MAR07050
	RRTSAM(IT)=0	MAR07060
	RRJ(IT)=0	MAR07070
	RRJM(IT)=0	MAR07080
	RRPEB(IT)=0	MAR07090
	RRPTEB(IT)=0	MAR07100
	RRPEBM(IT)=0	MAR07110
	RRPTBM(IT)=0	MAR07120
	RRJ1(IT)=0	MAR07130
	RRJ1M(IT)=0	MAR07140
C ***	RRJLT(IT)=0	MAR07150
	RRML(IT)=0	MAR07160
	KP=0	MAR07170
	DO 250 IM=1,NMOS	MAR07180
	DO 260 IL=1,NLOS	MAR07190
	DO 270 IG=1,NGRD	MAR07200

CC=CINV(IM,IL,IG,T+IT)	MAR07210
IF(CC.NE.0) THEN	MAR07220
AR=Y(IM,IL,IG,T+IT)/CC	MAR07230
ELSE	MAR07240
AR=0	MAR07250
ENDIF	MAR07260
IF(D(IM,IL,IG).NE.0.AND.PMLE(IM,IL,IG).NE.0	MAR07270
.AND.PMLE(IM,IL,IG).NE.1) THEN	MAR07280
KP=KP+1	MAR07290
RRML(IT)=RRML(IT)+(CC*(AR-PMLE(IM,IL,IG))**2)	MAR07300
/(PMLE(IM,IL,IG)*(1.-PMLE(IM,IL,IG)))	MAR07310
ENDIF	MAR07320
** PRODUCE CHI-SQUARE MEASURE OF EFFECTIVENESS	MAR07330
RRTSCA(IT)=RRTSCA(IT)+CHISQR(D(IM,IL,IG),TSCA(IM,IL,IG),	MAR07340
CC,AR)	MAR07350
RRTSAM(IT)=RRTSAM(IT)+CHISQR(D(IM,IL,IG),TSCAM(IM,IL,IG),	MAR07360
CC,AR)	MAR07370
RRJ (IT)=RRJ (IT)+CHISQR(D(IM,IL,IG),PJ (IM,IL,IG),	MAR07380
CC,AR)	MAR07390
RRJM (IT)=RRJM (IT)+CHISQR(D(IM,IL,IG),PJM (IM,IL,IG),	MAR07400
CC,AR)	MAR07410
RRJ1 (IT)=RRJ1 (IT)+CHISQR(D(IM,IL,IG),PJ1 (IM,IL,IG),	MAR07420
CC,AR)	MAR07430
RRJ1M (IT)=RRJ1M (IT)+CHISQR(D(IM,IL,IG),PJ1M(IM,IL,IG),	MAR07440
CC,AR)	MAR07450
RRPEB (IT)=RRPEB (IT)+CHISQR(D(IM,IL,IG),PEB (IM,IL,IG),	MAR07460
CC,AR)	MAR07470
RRPTEB (IT)=RRPTEB (IT)+CHISQR(D(IM,IL,IG),PTEB (IM,IL,IG),	MAR07480
CC,AR)	MAR07490
RRPEBM(IT)=RRPEBM(IT)+CHISQR(D(IM,IL,IG),PEBM(IM,IL,IG),	MAR07500
CC,AR)	MAR07510
RRPTBM(IT)=RRPTBM(IT)+CHISQR(D(IM,IL,IG),PTEBM(IM,IL,IG),	MAR07520
CC,AR)	MAR07530
C *** RRJLT (IT)=RRJLT (IT)+CHISQR(D(IM,IL,IG),PJLT(IM,IL,IG),	MAR07540
C CC,AR)	MAR07550
** PRODUCE MEAN ABSOLUTE DEVIATION MEASURE OF EFFECTIVENESS	MAR07560
IF(D(IM,IL,IG).NE.0.AND.PMLE(IM,IL,IG).NE.0	MAR07570
.AND.PMLE(IM,IL,IG).NE.1) THEN	MAR07580
CALL UOM(CC, AR, PMLE(IM,IL,IG), MOPMLE(IT),	MAR07590
MUPMLE(IT), KVOPML(IT), KVUPML(IT), KVDPML(IT))	MAR07600
ENDIF	MAR07610
IF(CC.NE.0) THEN	MAR07620
CALL UOM(CC, AR, TSCA(IM,IL,IG), MOTSCA(IT),	MAR07630
MUTSCA(IT), KVOTSA(IT), KVUTSA(IT), KVD TSA(IT))	MAR07640
CALL UOM(CC, AR, TSCAM(IM,IL,IG), MOTSAM(IT),	MAR07650
MUTSAM(IT), KVOTSM(IT), KVUTSM(IT), KVD TSM(IT))	MAR07660
CALL UOM(CC, AR, PJ(IM,IL,IG), MOPJ(IT),	MAR07670
MUPJ(IT), KVOPJ(IT), KVUPJ(IT), KVD PJ(IT))	MAR07680
CALL UOM(CC, AR, PJM(IM,IL,IG), MOPJM(IT),	MAR07690
MUPJM(IT), KVOPJM(IT), KVUPJM(IT), KVD PJM(IT))	MAR07700
CALL UOM(CC, AR, PJ1(IM,IL,IG), MOPJ1(IT),	MAR07710
	MAR07720
	MAR07730
	MAR07740
	MAR07750
	MAR07760

*	MUPJ1(IT), KVOPJ1(IT), KVUPJ1(IT), KVDPJ1(IT))	MAR07770
		MAR07780
	CALL UOM(CC, AR, PJ1M(IM,IL,IG), MOPJ1M(IT),	MAR07790
*	MUPJ1M(IT), KVOP1M(IT), KVUP1M(IT), KVDP1M(IT))	MAR07800
		MAR07810
	CALL UOM(CC, AR, PEB(IM,IL,IG), MOPEB(IT),	MAR07820
*	MUPEB(IT), KVOPEB(IT), KVUPEB(IT), KVDPEB(IT))	MAR07830
		MAR07840
	CALL UOM(CC, AR, PTEB(IM,IL,IG), MOPTEB(IT),	MAR07850
*	MUPTEB(IT), KVOPTB(IT), KVUPTB(IT), KVDPTB(IT))	MAR07860
		MAR07870
	CALL UOM(CC, AR, PEBM(IM,IL,IG), MOPEBM(IT),	MAR07880
*	MUPEBM(IT), KVOPBM(IT), KVUPBM(IT), KVDPBM(IT))	MAR07890
		MAR07900
	CALL UOM(CC, AR, PTEBM(IM,IL,IG), MOPTBM(IT),	MAR07910
*	MUPTBM(IT), KVOBTM(IT), KVUBTM(IT), KVDBTM(IT))	MAR07920
	ENDIF	MAR07930
		MAR07940
270	CONTINUE	MAR07950
260	CONTINUE	MAR07960
250	CONTINUE	MAR07970
	IF (KP .EQ. 0) THEN	MAR07980
	FF = 0	MAR07990
	ELSE	MAR08000
	FF=FLOAT(KV(IT))/FLOAT(KP)	MAR08010
	ENDIF	MAR08020
	RRML(IT)=RRML(IT)*FF	MAR08030
240	CONTINUE	MAR08040
**	COMPUTE TIME AVERAGE MAD	MAR08050
**	MEASURE OF EFFECTIVENESS	MAR08060
	DO 300 IT=1, NVYR	MAR08070
	MUPMLE(IT) = MUPMLE(IT)/KV(IT)	MAR08080
	MUTSCA(IT) = MUTSCA(IT)/KV(IT)	MAR08090
	MUTSAM(IT) = MUTSAM(IT)/KV(IT)	MAR08100
	MUPJ (IT) = MUPJ (IT)/KV(IT)	MAR08110
	MUPJM (IT) = MUPJM(IT)/KV(IT)	MAR08120
	MUPJ1 (IT) = MUPJ1(IT)/KV(IT)	MAR08130
	MUPJ1M(IT) = MUPJ1M(IT)/KV(IT)	MAR08140
	MUPEB (IT) = MUPEB(IT)/KV(IT)	MAR08150
	MUPTEB (IT) = MUPTEB(IT)/KV(IT)	MAR08160
	MUPEBM(IT) = MUPEBM(IT)/KV(IT)	MAR08170
	MUPTBM(IT) = MUPTBM(IT)/KV(IT)	MAR08180
	MOPMLE(IT) = MOPMLE(IT)/KV(IT)	MAR08190
	MOTSCA(IT) = MOTSCA(IT)/KV(IT)	MAR08200
	MOTSAM(IT) = MOTSAM(IT)/KV(IT)	MAR08210
	MOPJ (IT) = MOPJ (IT)/KV(IT)	MAR08220
	MOPJM (IT) = MOPJM(IT)/KV(IT)	MAR08230
	MOPJ1 (IT) = MOPJ1(IT)/KV(IT)	MAR08240
	MOPJ1M(IT) = MOPJ1M(IT)/KV(IT)	MAR08250
	MOPEB (IT) = MOPEB(IT)/KV(IT)	MAR08260
	MOPTEB (IT) = MOPTEB(IT)/KV(IT)	MAR08270
	MOPEBM(IT) = MOPEBM(IT)/KV(IT)	MAR08280
	MOPTBM(IT) = MOPTBM(IT)/KV(IT)	MAR08290
	MDPMLE(IT) = MUPMLE(IT) + MOPMLE(IT)	MAR08300
	MDTSCA(IT) = MUTSCA(IT) + MOTSCA(IT)	MAR08310
	MDTSAM(IT) = MUTSAM(IT) + MOTSAM(IT)	MAR08320

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MDPJ (IT) = MUPJ (IT) + MOPJ (IT)
MDPJM (IT) = MUPJM (IT) + MOPJM (IT)
MDPJ1 (IT) = MUPJ1 (IT) + MOPJ1 (IT)
MDPJ1M(IT) = MUPJ1M(IT) + MOPJ1M(IT)
MDPEB (IT) = MUPEB (IT) + MOPEB (IT)
MDPTEB (IT) = MUPTEB (IT) + MOPTEB (IT)
MDPEBM(IT) = MUPEBM(IT) + MOPEBM(IT)
MDPTBM(IT) = MUPTBM(IT) + MOPTBM(IT)
300 CONTINUE
*****
** PRINT SECTION FOR ALL OUTPUT. **
*****
** PRINT GROUP CHARACTERISTICS
WRITE(6, '(1X, 'TEN OR LESS BAYES VARIANCE ITERATIONS')')
WRITE(6, '(1X, I2, ' BAYES VARIANCE ITERATION(S)')') IC
WRITE(6, '(1X, 'XBB ', F12.6)') XBB
WRITE(6, 905) T, KK, SSB, SSE, SHR, SHR, ZB, ZTB, PZTB, A, AT
905 FORMAT( ' GROUP CHARACTERISTICS ', /, ' T=', I6, /, ' K=', I6,
* /, ' SSB=', F12.4, /, ' SSE=', F12.4, /, ' SHR=', F12.4,
* /, ' SHR=', F12.4, /, ' ZB=', F12.4, /, ' ZTB=', F12.4,
* /, ' PZTB=', F12.4, /, ' A=', F12.4, /, ' AT=', F12.4)
** PRINT RESULTS FOR TRANSFORM SPACE
WRITE(6, *) ' RESULTS FOR TRANSFORM SPACE'
WRITE(6, 906) (IVYR(I), I=1, NVYR)
906 FORMAT(14X, 10I12)
WRITE(6, 907) 'RTSCA ', (RTSCA(IT), IT=1, NVYR)
907 FORMAT(' ', A14, 10F12.4)
WRITE(6, 907) 'RJ ', (RJ (IT), IT=1, NVYR)
WRITE(6, 907) 'RJ1 ', (RJ1 (IT), IT=1, NVYR)
WRITE(6, 907) 'RXEB ', (RXEB (IT), IT=1, NVYR)
WRITE(6, 907) 'RXTEB ', (RXTEB (IT), IT=1, NVYR)
C *** WRITE(6, 907) 'RJLT ', (RJLT (IT), IT=1, NVYR)
WRITE(6, 907) 'RML ', (RML (IT), IT=1, NVYR)
C *** WRITE(6, 907) 'RSL ', (RSL (IT), IT=1, NVYR)
C *** WRITE(6, 907) 'RTJ ', (RTJ (IT), IT=1, NVYR)
C *** WRITE(6, 907) 'RTJLT ', (RTJLT (IT), IT=1, NVYR)
** PRINT RESULTS FOR ORIGINAL SPACE
WRITE(6, *) ' '
WRITE(6, *) ' RESULTS FOR ORIGINAL SPACE'
WRITE(6, 907) 'RRML ', (RRML (IT), IT=1, NVYR)
WRITE(6, 907) 'RRTSCA', (RRTSCA(IT), IT=1, NVYR)
WRITE(6, 907) 'RRTSAM', (RRTSAM(IT), IT=1, NVYR)
WRITE(6, 907) 'RRJ ', (RRJ (IT), IT=1, NVYR)
WRITE(6, 907) 'RRJM ', (RRJM (IT), IT=1, NVYR)
WRITE(6, 907) 'RRJ1 ', (RRJ1 (IT), IT=1, NVYR)
WRITE(6, 907) 'RRJ1M ', (RRJ1M (IT), IT=1, NVYR)
WRITE(6, 907) 'RRPEB ', (RRPEB (IT), IT=1, NVYR)
WRITE(6, 907) 'RRPTEB ', (RRPTEB (IT), IT=1, NVYR)
WRITE(6, 907) 'RRPEBM', (RRPEBM (IT), IT=1, NVYR)
WRITE(6, 907) 'RRPTBM', (RRPTBM (IT), IT=1, NVYR)
C *** WRITE(6, 907) 'RRJLT ', (RRJLT (IT), IT=1, NVYR)
WRITE(6, 908) 'KV ', (KV (IT), IT=1, NVYR)
908 FORMAT(' ', A14, 10I12)
WRITE(6, '(9X, 'KP ', I12)') KP
WRITE(6, *) ' '
** PRINT MAD CALCULATIONS

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WRITE(6, '(2X, 'PMLE')')	MAR08890
CALL WRMAD(IVYR, NVYR, MUPMLE, KVUPML, MOPMLE, KVOPML, MDPMLE, KVDPML)	MAR08900
WRITE(6, '(2X, 'TSCA')')	MAR08910
CALL WRMAD(IVYR, NVYR, MUTSCA, KVUTSA, MOTSCA, KVOTSA, MDTSCA, KVDTSA)	MAR08920
WRITE(6, '(2X, 'TSCAM')')	MAR08930
CALL WRMAD(IVYR, NVYR, MUTSAM, KVUTSM, MOTSAM, KVOTSM, MDTSAM, KVDTSM)	MAR08940
WRITE(6, '(2X, 'PJ')')	MAR08950
CALL WRMAD(IVYR, NVYR, MUPJ, KVUPJ, MOPJ, KVOPJ, MDPJ, KVDPJ)	MAR08960
WRITE(6, '(2X, 'PJM')')	MAR08970
CALL WRMAD(IVYR, NVYR, MUPJM, KVUPJM, MOPJM, KVOPJM, MDPJM, KVDPJM)	MAR08980
WRITE(6, '(2X, 'PJ1')')	MAR08990
CALL WRMAD(IVYR, NVYR, MUPJ1, KVUPJ1, MOPJ1, KVOPJ1, MDPJ1, KVDPJ1)	MAR09000
WRITE(6, '(2X, 'PJ1M')')	MAR09010
CALL WRMAD(IVYR, NVYR, MUPJ1M, KVUP1M, MOPJ1M, KVOP1M, MDPJ1M, KVDP1M)	MAR09020
WRITE(6, '(2X, 'PEB')')	MAR09030
CALL WRMAD(IVYR, NVYR, MUPEB, KVUPEB, MOPEB, KVOPEB, MDPEB, KVDPEB)	MAR09040
WRITE(6, '(2X, 'PEBM')')	MAR09050
CALL WRMAD(IVYR, NVYR, MUPEBM, KVUPBM, MOPEBM, KVOPBM, MDPEBM, KVDPBM)	MAR09060
WRITE(6, '(2X, 'PTEB')')	MAR09070
CALL WRMAD(IVYR, NVYR, MUPTB, KVUPTB, MOPTB, KVOPTB, MDPTB, KVDPB)	MAR09080
WRITE(6, '(2X, 'PTEBM')')	MAR09090
CALL WRMAD(IVYR, NVYR, MUPTBM, KVUBTM, MOPTBM, KVOBTM, MDPTBM, KVDBTM)	MAR09100
WRITE(6, *)	MAR09110
** PRINT THE CENTRAL INVENTORY MATRIX	MAR09120
WRITE(6, *) 'CINV MATRIX'	MAR09130
DO 1006 IT=1, NYR	MAR09140
WRITE(6, *) 'YEAR', IT	MAR09150
DO 1005 IL=1, NLOS	MAR09160
WRITE(6, 1007)(CINV(IM, IL, 1, IT), IM=1, NMOS)	MAR09170
1005 CONTINUE	MAR09180
1006 CONTINUE	MAR09190
1007 FORMAT(15(1X, F7.2))	MAR09200
WRITE(6, *)	MAR09210
** PRINT THE ATTRITION MATRIX	MAR09220
WRITE(6, *) 'Y MATRIX'	MAR09230
DO 1016 IT=1, NYR	MAR09240
WRITE(6, *) 'YEAR', IT	MAR09250
DO 1015 IL=1, NLOS	MAR09260
WRITE(6, 1007)(Y(IM, IL, 1, IT), IM=1, NMOS)	MAR09270
1015 CONTINUE	MAR09280
1016 CONTINUE	MAR09290
WRITE(6, *)	MAR09300
** PRINT ALL ARRAYS	MAR09310
WRITE(6, '(2X, 'ARRAY TSCA')')	MAR09320
CALL ARRAY(MXMOS, MXLOS, MXGRD, NMOS, NLOS, NGRD, TSCA)	MAR09330
WRITE(6, '(2X, 'ARRAY TSCAM')')	MAR09340
CALL ARRAY(MXMOS, MXLOS, MXGRD, NMOS, NLOS, NGRD, TSCAM)	MAR09350
WRITE(6, '(2X, 'ARRAY SHREB')')	MAR09360
CALL ARRAY(MXMOS, MXLOS, MXGRD, NMOS, NLOS, NGRD, SHREB)	MAR09370
WRITE(6, '(2X, 'ARRAY SHRTEB')')	MAR09380
CALL ARRAY(MXMOS, MXLOS, MXGRD, NMOS, NLOS, NGRD, SHRTEB)	MAR09390
WRITE(6, '(2X, 'ARRAY PMLE')')	MAR09400
CALL ARRAY(MXMOS, MXLOS, MXGRD, NMOS, NLOS, NGRD, PMLE)	MAR09410
WRITE(6, '(2X, 'ARRAY PJ')')	MAR09420
CALL ARRAY(MXMOS, MXLOS, MXGRD, NMOS, NLOS, NGRD, PJ)	MAR09430
WRITE(6, '(2X, 'ARRAY PJM')')	MAR09440

CALL ARRAY(MXMOS,MXLOS,MXGRD,NMOS,NLOS,NGRD,PJM)	MAR09450
WRITE(6,'(2X,'ARRAY PJ1')')	MAR09460
CALL ARRAY(MXMOS,MXLOS,MXGRD,NMOS,NLOS,NGRD,PJ1)	MAR09470
WRITE(6,'(2X,'ARRAY PJ1M')')	MAR09480
CALL ARRAY(MXMOS,MXLOS,MXGRD,NMOS,NLOS,NGRD,PJ1M)	MAR09490
WRITE(6,'(2X,'ARRAY PEB')')	MAR09500
CALL ARRAY(MXMOS,MXLOS,MXGRD,NMOS,NLOS,NGRD,PEB)	MAR09510
WRITE(6,'(2X,'ARRAY PTEB')')	MAR09520
CALL ARRAY(MXMOS,MXLOS,MXGRD,NMOS,NLOS,NGRD,PTTB)	MAR09530
WRITE(6,'(2X,'ARRAY PEBM')')	MAR09540
CALL ARRAY(MXMOS,MXLOS,MXGRD,NMOS,NLOS,NGRD,PEBM)	MAR09550
WRITE(6,'(2X,'ARRAY PTEBM')')	MAR09560
CALL ARRAY(MXMOS,MXLOS,MXGRD,NMOS,NLOS,NGRD,PTTB)	MAR09570
WRITE(6,'(2X,'ARRAY V')')	MAR09580
CALL ARRAY(MXMOS,MXLOS,MXGRD,NMOS,NLOS,NGRD,V)	MAR09590
WRITE(6,'(2X,'ARRAY XB')')	MAR09600
CALL ARRAY(MXMOS,MXLOS,MXGRD,NMOS,NLOS,NGRD,XB)	MAR09610
WRITE(6,'(2X,'ARRAY VT')')	MAR09620
CALL ARRAY(MXMOS,MXLOS,MXGRD,NMOS,NLOS,NGRD,VT)	MAR09630
WRITE(6,'(2X,'ARRAY XTB')')	MAR09640
CALL ARRAY(MXMOS,MXLOS,MXGRD,NMOS,NLOS,NGRD,XTB)	MAR09650
WRITE(6,'(2X,'ARRAY XJ')')	MAR09660
CALL ARRAY(MXMOS,MXLOS,MXGRD,NMOS,NLOS,NGRD,XJ)	MAR09670
WRITE(6,'(2X,'ARRAY XJ1')')	MAR09680
CALL ARRAY(MXMOS,MXLOS,MXGRD,NMOS,NLOS,NGRD,XJ1)	MAR09690
WRITE(6,'(2X,'ARRAY XEB')')	MAR09700
CALL ARRAY(MXMOS,MXLOS,MXGRD,NMOS,NLOS,NGRD,XEB)	MAR09710
WRITE(6,'(2X,'ARRAY XTEB')')	MAR09720
CALL ARRAY(MXMOS,MXLOS,MXGRD,NMOS,NLOS,NGRD,XTEB)	MAR09730
*	MAR09740
END	MAR09750
*****	MAR09760
SUBROUTINE WRMD(IYR,NYR,MU,KU,MO,KO,MD,KD)	MAR09770
*****	MAR09780
INTEGER IYR(NYR),KU(NYR),KO(NYR),KD(NYR)	MAR09790
REAL MU(NYR),MO(NYR),MD(NYR)	MAR09800
WRITE(6,'(/9X,'DEVIATION',T26,I2,T42,I2,T59,I2/)'')	MAR09810
* (IYR(IW),IW=1,NYR)	MAR09820
WRITE(6,'(9X,'UNDERAGE',T22,F9.6,'('',I2,'')',T38,F9.6,	MAR09830
* '(',I2,'')',T55,F9.6,'('',I2,'')'')	MAR09840
* MU(1),KU(1),MU(2),KU(2),MU(3),KU(3)	MAR09850
WRITE(6,'(9X,'OVERAGE',T22,F9.6,'('',I2,'')',T38,F9.6,	MAR09860
* '(',I2,'')',T55,F9.6,'('',I2,'')'')	MAR09870
* MO(1),KO(1),MO(2),KO(2),MO(3),KO(3)	MAR09880
WRITE(6,'(9X,'TOTAL',T22,F9.6,'('',I2,'')',T38,F9.6,	MAR09890
* '(',I2,'')',T55,F9.6,'('',I2,'')'')	MAR09900
* MD(1),KD(1),MD(2),KD(2),MD(3),KD(3)	MAR09910
WRITE(6,'(/)'')	MAR09920
RETURN	MAR09930
END	MAR09940
*****	MAR09950
SUBROUTINE ARRAY(MM,ML,MG,NM,NL,NG,OBJ1)	MAR09960
*****	MAR09970
REAL OBJ1(MM,ML,MG)	MAR09980
DO 10 IG1=1,NG	MAR09990

DO 10 IL1=1,NL	MAR10000
10 WRITE(6,'(2X,15(F7.3,1X))') (OBJ1(IM1,IL1,IG1),IM1=1,NM)	MAR10010
WRITE(6,'(//)')	MAR10020
RETURN	MAR10030
END	MAR10040
*****	MAR10050
SUBROUTINE FTT(CINV,Y, X)	MAR10060
*****	MAR10070
REAL CINV,X,Y	MAR10080
TEMP=Y/(1+CINV)	MAR10090
TEMP1=(1.+Y)/(1+CINV)	MAR10100
TEMP=SQRT(.5+CINV)*.5*(ASIN(-1.+2.*TEMP) +	MAR10110
* ASIN(-1.+2.*TEMP1))	MAR10120
X=TEMP	MAR10130
END	MAR10140
*****	MAR10150
SUBROUTINE UOM(VINV,ACT,OBJ,OVER,UNDER,KO,KU,KD)	MAR10160
*****	MAR10170
REAL OBJ	MAR10180
IF (ACT .GT. OBJ) THEN	MAR10190
UNDER = UNDER+VINV*(ACT-OBJ)	MAR10200
KU=KU+1	MAR10210
ELSE	MAR10220
OVER = OVER +VINV*(OBJ-ACT)	MAR10230
KO=KO+1	MAR10240
ENDIF	MAR10250
KD=KD+1	MAR10260
RETURN	MAR10270
END	MAR10280
*****	MAR10290
REAL FUNCTION RHO(XB,XBB,SSE,SSB,DEE,KK,T,NN)	MAR10300
*****	MAR10310
INTEGER T	MAR10320
REAL*8 SIGSQ, V, U	MAR10330
SIGSQ=SSE/(NN+2)	MAR10340
V=SSB/T	MAR10350
U=KK*(KK-3)*(XB-XBB)**2/(V*(KK-1))	MAR10360
RHO=DMIN1(1.DO, DEE/DSQRT(U))	MAR10370
END	MAR10380
*****	MAR10390
REAL FUNCTION MILLER(X, AVGINV)	MAR10400
*****	MAR10410
TM = X/SQRT(AVGINV+.5)+1.570796	MAR10420
TEMP1 = ASIN(SQRT(1./(AVGINV+1)))	MAR10430
IF (TM .LT. TEMP1) THEN	MAR10440
MILLER = 0.0	MAR10450
GOTO 1	MAR10460
ENDIF	MAR10470
IF (TM .GT. 3.14159-TEMP1) THEN	MAR10480
MILLER = 1.0	MAR10490
GOTO 1	MAR10500
ENDIF	MAR10510
IF (COS(TM) .EQ. 0.0) THEN	MAR10520
MILLER = 0.5	MAR10530
GOTO 1	MAR10540
ENDIF	MAR10550

IF (COS(TM) .GT. 0.0) SIGNUM = 1.	MAR10560
IF (COS(TM) .LT. 0.0) SIGNUM = -1.	MAR10570
TEMP2=1-(SIN(TM)+(SIN(TM)-(1./SIN(TM)))/AVGINV)**2	MAR10580
IF (TEMP2.LT.0.0) TEMP2=0.0	MAR10590
MILLER=.5*(1.-SIGNUM*TEMP2**.5)	MAR10600
1 END	MAR10610
*****	MAR10620
REAL FUNCTION SCINV(X,A)	MAR10630
*****	MAR10640
** SCALE INVERSION (TRANSFORM ==> ORIGINAL SCALE)	MAR10650
R=X/SQRT(A+.5)	MAR10660
IF(R .LT. -1.570796) THEN	MAR10670
SCINV=0.	MAR10680
ELSE IF(R .GT. 1.570796) THEN	MAR10690
SCINV=1.	MAR10700
ELSE	MAR10710
SCINV=.5*(1.+SIN(R))	MAR10720
ENDIF	MAR10730
END	MAR10740
*****	MAR10750
REAL FUNCTION VAR(ZZ)	MAR10760
*****	MAR10770
DATA AA/1.6835/, B1/-.8934/, B2/.8991/	MAR10780
IF (ZZ .GT. 2.2) THEN	MAR10790
VAR = 1.	MAR10800
ELSE	MAR10810
IF (ZZ .LT. 1.001) THEN	MAR10820
VAR = .05	MAR10830
ELSE	MAR10840
VAR = AA*(ZZ**B1)*(ZZ-1.)**B2	MAR10850
ENDIF	MAR10860
ENDIF	MAR10870
END	MAR10880
*****	MAR10890
REAL FUNCTION CHISQR(ID,OBJ,INV,ACT)	MAR10900
*****	MAR10910
REAL OBJ,INV	MAR10920
IF (ID.NE.0 .AND. OBJ.NE.0. .AND. OBJ.NE.1.)	MAR10930
* CHISQR = (INV*(ACT-OBJ)**2)/(OBJ*(1.-OBJ))	MAR10940
END	MAR10950
*****	MAR10960
SUBROUTINE READER(MOS,LOS,GRADE,YR,INV,Y, NMOS,SMOS,	MAR10970
* NGRD,SGRD, SLOS1,SLOS2, ST1,ST2, IM,IL,IG,IT, IEOF)	MAR10980
*****	MAR10990
** READ RECORD AND ACCUMULATE LOSSES	MAR11000
INTEGER MOS, LOS, GRADE, YR, INV, LOSS(8)	MAR11010
REAL Y	MAR11020
INTEGER ST1, ST2	MAR11030
INTEGER SLOS1, SLOS2	MAR11040
INTEGER SMOS(40), NMOS	MAR11050
INTEGER SGRD(10), NGRD	MAR11060
ICNT=0	MAR11070
1 READ(1,100,END=999) YR,MOS,GRADE,LOS,INV,(LOSS(I),I=1,8)	MAR11080
100 FORMAT(4I2,9I5)	MAR11090
** CHECK IF RECORD MEETS SELECTION CRITERIA. OTHERWISE REJECT.	MAR11100
IM=0	MAR11110

	DO 10 I=1,NMOS	MAR11120
	IF(MOS .EQ. SMOS(I)) THEN	MAR11130
	IM=I	MAR11140
	GO TO 20	MAR11150
	END IF	MAR11160
10	CONTINUE	MAR11170
	GO TO 1	MAR11180
*		MAR11190
20	CONTINUE	MAR11200
	DO 30 I=1,NGRD	MAR11210
	IF(GRADE .EQ. SGRD(I)) THEN	MAR11220
	IG=I	MAR11230
	GO TO 40	MAR11240
	END IF	MAR11250
30	CONTINUE	MAR11260
	GO TO 1	MAR11270
40	CONTINUE	MAR11280
*		MAR11290
	IF(LOS .LT. SLOS1 .OR. LOS .GT. SLOS2) GO TO 1	MAR11300
	IL=LOS-SLOS1+1	MAR11310
**	YEARS OVER ST2 ARE USED FOR VALIDATION	MAR11320
	IF(YR .LT. ST1) GO TO 1	MAR11330
	IT=YR-ST1+1	MAR11340
**	COMPUTE TOTAL LOSS	MAR11350
	Y=0	MAR11360
	DO 50 I=1,8	MAR11370
	Y=Y + LOSS(I)	MAR11380
50	CONTINUE	MAR11390
*		MAR11400
	RETURN	MAR11410
999	IEOF=1	MAR11420
	END	MAR11430

APPENDIX C. SIMULATION FOR CHOICE OF AVERAGE INVENTORY VALUES

```

PROGRAM MCSIM
*****
* SIMULATION COMPARISON OF INVENTORY MEAN PERFORMANCE *
*
* PURPOSE: TO IDENTIFY THE BEST MEAN VALUE (ARITHMETIC, *
*          GEOMETRIC AND HARMONIC) FOR USE IN THE INVERSION OF *
*          TRANSFORMED DATA TO THE ORIGINAL SPACE. *
* PROGRAMMER: CAPT C R DICKINSON, USMC *
* VARIABLE DESCRIPTION: *
*          LAMBDA - PARAMETER USED IN THE POISSON RANDOM NUMBER *
*          GENERATOR *
*          TRIALS - VECTOR OF INVENTORY VALUES GENERATED FROM *
*          THE POISSON RANDOM NUMBER GENERATOR *
*          REPS - NUMBER OF INVENTORY VALUES TO BE GENERATED *
*          BASIC - BASIC INVERSION OF TRANSFORMED VALUES *
*          FTE1 - FREEMAN-TUKEY EXACT INVERSION WITH *
*          *
*          FTE2 - FREEMAN-TUKEY EXACT INVERSION WITH *
*          *
*          Z - TRANSFORMED SCALE VARIABLE *
*          MEANS - VECTOR OF ARITHMETIC, GEOMETRIC AND HARMONIC *
*          MEANS *
*          TRSUM - SUMMATION OF TRIALS FOR USE IN THE *
*          ARITHMETIC MEAN *
*          TRMULT - MULTIPLICATION OF TRIALS FOR USE IN THE *
*          GEOMETRIC MEAN *
*          TRINV - SUMMATION OF THE INVERSE OF TRIALS FOR USE *
*          IN THE HARMONIC MEAN *
* INPUT/OUTPUT: *
*          INPUT - HARD CODED PARAMETERS *
*          OUTPUT - SIM OUT A1 (DISK FILE 03) *
*****
*
*** VARIABLE DECLARATION
INTEGER TRIALS(100),Y(100),REPS(3),PROB,R,
C COUNT,TRSUM,YTEMP,YSUM
REAL Z(100),T(100),MEAN(3),BASIC(3),DIFF(3),LAMBDA(5),
C P(4),FTE1(3),FTE2(3),TRMULT
DOUBLE PRECISION DPSEED
*
CALL EXCMS('FILEDEF 02 DISK REPLIC LISTING A1')
*
*** VARIABLE INITIALIZATION
PI=3.14159
DPSEED=889246.D0
LAM=5
R=3
PROB=4
COUNT=2

```

```

MCS00010
MCS00020
MCS00030
MCS00040
MCS00050
MCS00060
MCS00070
MCS00080
MCS00090
MCS00100
MCS00110
MCS00120
MCS00130
MCS00140
MCS00150
MCS00160
MCS00170
MCS00180
MCS00190
MCS00200
MCS00210
MCS00220
MCS00230
MCS00240
MCS00250
MCS00260
MCS00270
MCS00280
MCS00290
MCS00300
MCS00310
MCS00320
MCS00330
MCS00340
MCS00350
MCS00360
MCS00370
MCS00380
MCS00390
MCS00400
MCS00410
MCS00420
MCS00430
MCS00440
MCS00450
MCS00460
MCS00470
MCS00480
MCS00490

```

	DATA REPS/20,40,60/	MCS00500
	DATA LAMBDA/1.,2.,4.,8.,16./	MCS00510
	DATA P/.05,.1,.2,.4/	MCS00520
*		MCS00530
***	HEADER PRINT	MCS00540
	WRITE (02, '(/T6, 'LAMBDA', T15, 'REPS', T27, 'P',	MCS00550
C	T34, 'METHOD', T46, 'ARITHMETIC', T59,	MCS00560
C	'GEOMETRIC', T71, 'HARMONIC' /))	MCS00570
*		MCS00580
***	LOOP FOR LAMBDA PARAMETER	MCS00590
	DO 500 I=1, LAM	MCS00600
	DO 450 J=1, R	MCS00610
	CALL GGPOS(LAMBDA(I), DPSEED, REPS(J), TRIALS, IER)	MCS00620
	NSUM=0	MCS00630
	DO 10 J1=1, REPS(J)	MCS00640
	TRIALS(J1)=TRIALS(J1)+1	MCS00650
10	NSUM=NSUM+TRIALS(J1)	MCS00660
*		MCS00670
***	LOOP FOR PROBABILITY PARAMETER	MCS00680
	DO 400 K=1, PROB	MCS00690
	COUNT=COUNT+4	MCS00700
***	LINE COUNTER FOR PAGE BREAK	MCS00710
	IF (COUNT.GT.55) THEN	MCS00720
	WRITE (2, '(1H1)')	MCS00730
	COUNT=2	MCS00740
	WRITE (02, '(/T6, 'LAMBDA', T15, 'REPS', T27, 'P',	MCS00750
C	T34, 'METHOD', T46, 'ARITHMETIC', T59,	MCS00760
C	'GEOMETRIC', T71, 'HARMONIC' /))	MCS00770
	ENDIF	MCS00780
*		MCS00790
	TRSUM=0	MCS00800
	TRINV=0	MCS00810
	TRMULT=0.	MCS00820
	ARCSUM=0.	MCS00830
	TSUM=0.	MCS00840
	YSUM=0	MCS00850
*		MCS00860
***	LOOP FOR REPITITION PARAMETER	MCS00870
	DO 20 K1=1, REPS(J)	MCS00880
	CALL GGBN(DPSEED, 1, TRIALS(K1), P(K), YTEMP)	MCS00890
	YSUM=YSUM+YTEMP	MCS00900
	TRSUM=TRSUM+TRIALS(K1)	MCS00910
	TRMULT=TRMULT+ALOG(REAL(TRIALS(K1)))	MCS00920
	TRINV=TRINV+1./REAL(TRIALS(K1))	MCS00930
	Y(K1)=YTEMP	MCS00940
	Z(K1)=.5*(TRIALS(K1)+.5)**.5*(ASIN(2.*Y(K1)/	MCS00950
C	(TRIALS(K1)+1.)-1.)+ASIN(2.*(Y(K1)+1.)/	MCS00960
C	(TRIALS(K1)+1.)-1.)+PI)	MCS00970
	T(K1)=Z(K1)/((TRIALS(K1)+.5)**.5)	MCS00980
	ARCSUM=ARCSUM+Z(K1)	MCS00990
20	TSUM=TSUM+T(K1)	MCS01000
*		MCS01010
	ZAVG=ARCSUM/REAL(REPS(J))	MCS01020
	TAVG=TSUM/REAL(REPS(J))	MCS01030
***	MEAN VALUE CALCULATION	MCS01040
		MCS01050

	MEAN(1)=TRSUM/REAL(REPS(J))	MCS01060
	MEAN(2)=EXP(TRMULT*(1./REAL(REPS(J))))	MCS01070
	MEAN(3)=REAL(REPS(J))/TRINV	MCS01080
*		MCS01090
	IF (COS(TAVG).GE.0.0) THEN	MCS01100
	SIGNUM=1.	MCS01110
	ELSE	MCS01120
	SIGNUM=-1.	MCS01130
	ENDIF	MCS01140
*		MCS01150
	DO 30 K2=1,3	MCS01160
	TAVG1=ZAVG/(MEAN(K2)+.5)**.5	MCS01170
	IF (COS(TAVG1).GE.0.0) THEN	MCS01180
	SIG1=1.	MCS01190
	ELSE	MCS01200
	SIG1=-1.	MCS01210
	ENDIF	MCS01220
***	BASIC AND FTE INVERSE METHOD CALCULATION	MCS01230
	FTE1(K2)=.5*(1.-(SIGNUM*(1.-(ABS(SIN(TAVG)+	MCS01240
C	(SIN(TAVG)-(1./SIN(TAVG)))/MEAN(K2))**2.))**5))	MCS01250
*		MCS01260
	FTE2(K2)=.5*(1.-(SIG1*(1.-(ABS(SIN(TAVG1)+	MCS01270
C	(SIN(TAVG1)-(1./SIN(TAVG1)))/MEAN(K2))**2.))**5))	MCS01280
*		MCS01290
30	BASIC(K2)=.5*(1.+SIN(ZAVG/((MEAN(K2)+.5)**.5)	MCS01300
C	-(PI/2.)))	MCS01310
***	OUTPUT STATEMENTS	MCS01320
	DO 40 K3=1,3	MCS01330
	IF (K3.EQ.1) THEN	MCS01340
	WRITE (02,'(T6,F4.1,T15,I3,T25,F5.3,T35,'BASIC',	MCS01350
C	T46,F7.4,T59,F7.4,T71,F7.4)')	MCS01360
C	LAMBDA(I),REPS(J),P(K),(BASIC(IK),IK=1,3)	MCS01370
	ELSE	MCS01380
	IF (K3.EQ.2) THEN	MCS01390
	WRITE (02,'(T35,'FTE1',T46,F7.4,T59,F7.4,	MCS01400
C	T71,F7.4)') (FTE1(IK),IK=1,3)	MCS01410
	ELSE	MCS01420
	WRITE (02,'(T35,'FTE2',T46,F7.4,T59,F7.4,	MCS01430
C	T71,F7.4)') (FTE2(IK),IK=1,3)	MCS01440
	ENDIF	MCS01450
	ENDIF	MCS01460
40	CONTINUE	MCS01470
*		MCS01480
400	CONTINUE	MCS01490
450	CONTINUE	MCS01500
500	CONTINUE	MCS01510
	STOP	MCS01520
	END	MCS01530

LIST OF REFERENCES

1. Tucker, D.E., Loss Rate Estimation in Marine Corps Officer Manpower Models, Master's Thesis, Naval Postgraduate School, Monterey, California, September 1985.
2. Robinson, J.R., Limited Translation Shrinkage Estimation of Loss Rates in Marine Corps Manpower Models, Master's Thesis, Naval Postgraduate School, Monterey, California, March 1986.
3. Larsen, R.W., The Aggregation of Population Groups to Improve the Predictability of Marine Corps Officer Attrition Estimation, Master's Thesis, Naval Postgraduate School, Monterey, California, December 1987.
4. Carter, G.M. and Rolph, J.E., "Empirical Bayes Methods Applied to Estimating Fire Alarm Probabilities", Journal of the American Statistical Association, v. 69, no. 248, pp. 880-885, December 1974.
5. Miller, J.J., "The Inversion of the Freeman-Tukey Double Arcsine Transformation", The American Statistician, v. 32, no. 4, p. 138, November 1978.
6. Casella, G., "An Introduction to Empirical Bayes Data Analysis", The American Statistician, v. 39, no. 2, pp. 83-87, May 1985.
7. Bres, E.S. and Rowe, M.W., "Development and Analysis of Loss Rate Forecasting Techniques for the Navy's Unrestricted Line (URL) Officers", Navy Personnel Research and Development Center Technical Report 79-20, June 1979.
8. Larson, H.J., Introduction to Probability Theory and Statistical Inference, John Wiley and Sons, Inc., 1982.
9. Efron, B. and Morris C., "Stein Estimation Rule and Its Competitors--An Empirical Bayes Approach", Journal of the American Statistical Association, v. 68, pp. 117-130, 1973.

10. Fay, R.E. III and Herriot, R.A., "Estimation of Income for Small Places: An Application of James-Stein Procedures to Census Data", Journal of the American Statistical Association, V. 74, no. 366, pp. 269-277, June 1979.
11. Stein, C., "Inadmissibility of the Usual Estimator for the Mean of a Multivariate Normal Distribution", Proceedings of the Third Berkeley Symposium of Mathematical Statistics and Probability, v. 1, Berkeley: University of California Press, pp. 197-206, 1955.

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